

## SOLUTIONS to Quiz #1

### Problem 1 (20%)

Define  $\xi = (X - \mu_X) / \sigma_X$  and  $\eta = (Y - \mu_Y) / \sigma_Y$ , then we have the pdf

$$f(\xi, \eta) = \frac{1}{(\sqrt{2\pi})^2 \sqrt{1 - \rho^2}} \exp\left[-\frac{1}{2(1 - \rho^2)}(\xi^2 + \eta^2 - 2\rho\xi\eta)\right]$$

$$\begin{aligned} \text{(a) } f(\eta) &= \frac{1}{(\sqrt{2\pi})} * \frac{1}{\sqrt{2\pi(1 - \rho^2)}} \int_{-\infty}^{\infty} \exp\left\{-\frac{1}{2(1 - \rho^2)}[(\xi - \rho\eta)]^2 + \eta^2 - \rho^2\eta^2\right\} d\xi \\ &= \frac{1}{(\sqrt{2\pi})} \exp\left[-\frac{\eta^2}{2}\right] \Rightarrow \eta \text{ is a Gaussian random variable} \end{aligned}$$

so the pdf of  $X$  is  $f_X(x) = \frac{1}{\sqrt{2\pi}\sigma_X} \exp\left[-\frac{(x - \mu_X)^2}{2\sigma_X^2}\right]$ ;

$$\text{(b) } f(\xi | \eta) = f(\xi, \eta) / f(\eta) = \frac{1}{\sqrt{2\pi(1 - \rho^2)}} \exp\left\{-\frac{1}{2(1 - \rho^2)}[(\xi - \rho\eta)]^2\right\}$$

or the pdf of  $X$  given  $Y = y$  is Gaussian with the following pdf:

$$f(x | y) = \frac{1}{\sqrt{2\pi(1 - \rho^2)}\sigma_X} \exp\left\{-\frac{1}{2(1 - \rho^2)\sigma_X^2}\left[(x - \mu_X) - \frac{\rho\sigma_X}{\sigma_Y}(y - \mu_Y)\right]^2\right\};$$

(c) The conditional variance of  $X$  given  $Y = y$  is  $(1 - \rho^2)\sigma_X^2$ , and the

conditional mean of  $X$  given  $y$  is  $\mu_X + \frac{\rho\sigma_X}{\sigma_Y}(y - \mu_Y)$ ;

(d) Since  $(1 - \rho^2)\sigma_X^2 \leq \sigma_X^2$ , knowing  $Y = y$  will reduce the variance of  $X$  given  $y$ .

### Problem 2 (20%)

$$\mu_X = \mu_Y = \mu_Z = 0, \sigma_X^2 = \sigma_Y^2 = \sigma_Z^2 = 1, \text{ and } \rho_{XY} = 0, \rho_{XZ} = 1/2, \rho_{YZ} = -1/2;$$

$$W = X + Y + Z, \text{ and } U = Y - Z \Rightarrow \mu_W = \mu_U = 0.$$

$$\begin{aligned} \text{(a) } \sigma_W^2 &= E[(W - \mu_W)^2] = E[X^2 + Y^2 + Z^2 + 2XY + 2YZ + 2XZ] \\ &= \sigma_X^2 + \sigma_Y^2 + \sigma_Z^2 + 2\rho_{XY}\sigma_X\sigma_Y + 2\rho_{YZ}\sigma_Y\sigma_Z + 2\rho_{XZ}\sigma_X\sigma_Z = 3; \end{aligned}$$

$$\begin{aligned} \text{(b) } \rho_{WX} &= E[(W - \mu_W)(X - \mu_X)] / \sigma_W\sigma_X = E[X^2 + XY + XZ] / \sigma_W\sigma_X \\ &= [\sigma_X^2 + \rho_{XY}\sigma_X\sigma_Y + \rho_{XZ}\sigma_X\sigma_Z] / \sigma_W\sigma_X = \sqrt{3}/2; \end{aligned}$$

$$\text{(c) } \sigma_U^2 = E[(U - \mu_U)^2] = E[Y^2 - 2YZ + Z^2] = 3$$

$$\begin{aligned} \rho_{WU} &= E[(W - \mu_W)(U - \mu_U)] / \sigma_W\sigma_U = E[XY - XZ + Y^2 - Z^2] / \sigma_W\sigma_U \\ &= (\rho_{XY} - \rho_{XZ} + \sigma_Y^2 - \sigma_Z^2) / \sigma_W\sigma_U = (-\frac{1}{2}) / 3 = -1/6. \end{aligned}$$

### Problem 3 (20%)

$X(t) = \cos(100t + \Theta)$  and  $Y(t) = \cos(100t + \Psi)$ , with both  $\Theta$  and  $\Psi$  being independent, and uniformly distributed from 0 to  $2\pi$ .

(a)  $Z(t) = X(t) + Y(t) = \cos(100t + \Theta) + \cos(100t + \Psi) = A \cos(100t + \Phi)$ ; with  
 $A = 2 \cos\left(\frac{\Theta - \Psi}{2}\right)$  and  $\Phi = \frac{\Theta + \Psi}{2}$  [ $\cos X + \cos Y = 2 \cos\left(\frac{X - Y}{2}\right) \cos\left(\frac{X + Y}{2}\right)$ ];

(b)  $\Phi$  has a triangular distribution between 0 and  $2\pi$  with a pdf

$$f_{\Phi}(\phi) = \begin{cases} \phi / \pi^2 & 0 \leq \phi \leq \pi \\ (2\pi - \phi) / \pi^2 & \pi \leq \phi \leq 2\pi; \\ 0 & \text{otherwise} \end{cases}$$

$$\text{so } P(\phi \leq \pi/3) = \int_0^{\pi/3} f_{\Phi}(\phi) d\phi = \phi^2 / 2\pi^2 \Big|_0^{\pi/3} = 1/18;$$

(c) Define  $\alpha = (\Theta - \Psi) / 2$ , then  $\alpha$  has a triangular distribution between  $-\pi$  and  $\pi$  with

$$f(\alpha) = \begin{cases} (\phi + \pi) / \pi^2 & -\pi \leq \phi \leq 0 \\ (\pi - \phi) / \pi^2 & 0 \leq \phi \leq \pi; \\ 0 & \text{otherwise} \end{cases}$$

(d)  $P(A > 1) = P(\cos \alpha > 1/2) = P(|\alpha| < \pi/3)$

$$= \int_{-\pi/3}^{\pi/3} f(\alpha) d\alpha = -2(\pi - \phi)^2 / 2\pi^2 \Big|_0^{\pi/3} = -(4/9) + 1 = 5/9.$$

#### Problem 4 (20%)

(a) The total number of possible combinations for drawing two cards is

$J = {}_{52}C_2 = (52 * 51) / 2 = 1326$ . There are 4 Aces and 16 cards (10, Jack, Queen and King, four each) to make up a total of  $K = 64$  combinations of blackjack hands. So the probability of  $K / J = 64 / 1326 = 0.04826546$ ;

(b) Let  $A$  be the event that the dealer get a blackjack hand and  $B$  be the event that the player get a blackjack hand, then  $P(A \text{ and } B) = P(A | B)P(B) = 0.01773$ ;

[since  $P(A | B) = [(4 - 1) * (16 - 1)] / {}_{50}C_2 = 45 / 1225 = 0.036735$ ,  $P(B) = 0.048265$ ]

(c) The total number of possible combinations for drawing five cards is

$M = {}_{52}C_5 = (52 * 51 * 50 * 49 * 48) / (5 * 4 * 3 * 2) = 2598960$ . To get a full house hand, there are  $13 * 12$  types of three-of-a-kind and pair combinations, so the total number is  $N = 13 * 12 * {}_4C_3 * {}_4C_2 = 3744$ . So  $P(\text{"Full House"}) = N / M = 0.001441$ ;

(d) Continuing from parts (c), to get a four-of-a-kind hand, there are  $L = 13 * 48 = 624$  combinations, so  $P(\text{"four-of-a-kind"}) = L / M = 0.0002401$ .

#### Problem 5 (10%)

(a)  $P(\text{R appears}) = P(\text{R|R})P(\text{R}) + P(\text{R|E})P(\text{E}) + P(\text{R|T})P(\text{T})$

$$= 0.96 * 0.0484 + 0.02 * 0.1031 + 0.02 * 0.0796 = 0.0501;$$

(b)  $P(\text{Error | R appears}) = [P(\text{R|E})P(\text{E}) + P(\text{R|T})P(\text{T})] / P(\text{R appears})$

$$= [0.02 * 0.1031 + 0.02 * 0.0796] / 0.0501 = 0.0729.$$

**Problem 6 (20%)**

(a)  $P_n(X = k) = {}_n C_k p^k (1 - p)^{n-k}$

$$\begin{aligned} E[X] &= \sum_{k=0}^n \frac{n!}{k!(n-k)!} k p^k (1-p)^{n-k} \\ &= np * \sum_{k=1}^n \frac{(n-1)!}{(k-1)!(n-k)!} p^{k-1} (1-p)^{n-k} \\ &= np * \sum_{l=0}^{n-1} \frac{l!}{l!(n-l)!} p^l (1-p)^{n-l} = np \end{aligned}$$

[with a change of variable of  $m = n - 1$  and  $l = k - 1$ ]

$$\begin{aligned} E[X^2 - X] &= \sum_{k=0}^n \frac{n!}{k!(n-k)!} (k^2 - k) p^k (1-p)^{n-k} \\ &= n(n-1)p^2 * \sum_{k=2}^n \frac{(n-2)!}{(k-2)!(n-k)!} p^{k-2} (1-p)^{n-k} \\ &= n(n-1)p^2 * \sum_{l=0}^{n-2} \frac{l!}{l!(n-l)!} p^l (1-p)^{n-l} = n(n-1)p^2 \end{aligned}$$

$$\Rightarrow \text{Var}[X] = E[X^2 - X] + E[X] - (E[X])^2 = n(n-1)p^2 + np - n^2 p^2 = np(1-p);$$

(b)  $\phi(u) = (1 - p + pe^{ju})^n$  and  $E[X^m] = j^{-m} * d^m \phi(u) / du^m |_{u=0}$ .

Since  $\frac{d\phi(u)}{du} = npje^{ju} (1 - p + pe^{ju})^{n-1}$  and

$$\frac{d^2\phi(u)}{du^2} = npj^2 e^{ju} (1 - p + pe^{ju})^{n-1} + n(n-1)pje^{ju} (1 - p + pe^{ju})^{n-2} * pje^{ju}$$

so  $E[X] = j^{-1} d\phi(u) / du |_{u=0} = np$  [same as in part (a)],

$$E[X^2] = j^{-2} d^2\phi(u) / du^2 |_{u=0} = np + n(n-1)p^2 = (np)^2 + np(1-p),$$

therefore the variance  $\text{Var}(X) = E[X^2] - (E[X])^2 = np(1-p)$  [same as in part (a)];

(c) (*Extra Credit +10%*) For  $n = 100,000$ , and  $p = 0.5$ , we have  $\mu = E[X] = 50,000$ ,

and  $\sigma = \sqrt{\text{Var}(X)} = \sqrt{25000} = 158.11$ . From the given table, we know that

with a Gaussian approximation  $P(\mu - 3\sigma < X < \mu + 3\sigma) = 0.9974$ , or

$P(49525.67 < X < 50474.33) = 0.9974$  if the voters do not favor either

candidate. The actual value of  $X = 49,000$  shows that the probability that

Anderson will win is almost zero, and therefore his opponent is likely to win.