

Through the Ideology of the Beholder: How Ideology Shapes Perceptions of Partisan Groups*

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Abstract

Growing attitudinal and affective differences across party lines and increasing social polarization are often attributed to the strengthening of partisanship as a social identity. Scholars have paid less attention to personal preferences as a contributor to these phenomena. Our focus is on how citizens' policy beliefs—their operational ideologies—are associated with their views of partisan groups. We examine our perspective with two studies. In the first, we find that the attribution of ideologically extreme political views to an individual's peer significantly reduces interest in interpersonal interaction but find limited evidence that partisan group membership alone induces social polarization. In the second, we show that citizens' policy views are strongly associated with their perceptions of their own partisan group as well as their counterpartisans. Together, our results have important implications for understanding the consequences of increased polarization and partisan antipathy in contemporary politics.

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Partisans in the United States are increasingly divided not just by their views of political candidates but also in how they view each other. Over the last several decades, Democrats have registered increasingly negative views of Republicans just as Republicans have expressed increasingly negative views of Democrats (e.g., Abramowitz and Webster 2016; Iyengar, Sood, and Lelkes 2012). By and large, these trends have been characterized as affective in nature, driven by the “power of partisanship as a social identity” (Iyengar et al. 2019, 130) and where voters base their partisan loyalties more on fear and loathing of the out-party than in-group affinity (Abramowitz and Webster 2016, 2018).

In this paper, we study how Americans view each other across party lines and consider the connection between citizens’ policy views and their views of partisan groups. Though partisan identity remains important, our perspective suggests that existing scholarship generally fails to appreciate the extent of ideology’s role in shaping mass attitudes and behaviors. While studies at the elite level show that increased ideological extremism in Congress is associated with more partisan rhetoric (e.g., Gentzkow, Shapiro, and Taddy 2016), less intraparty collaboration (e.g., Desmarais et al. 2015), and more distant social ties (e.g., Alduncin, Parker, and Theriault 2017), at the mass level researchers more commonly emphasize how group attachments and identities (e.g. Iyengar, Sood, and Lelkes 2012; Iyengar and Westwood 2015), psychological traits (e.g., Simas, Clifford, and Kirkland 2020), and partisan-ideological sorting (e.g. Levendusky 2009; Mason 2015) produce and reinforce political and social polarization.

We argue that evaluations of partisan out-group members are driven primarily by the *combination* of partisanship and ideological considerations. We develop two primary theoretical expectations. First, we argue that social evaluations reflect differences in ideological positions rather than reflexive responses on the basis of partisanship. Our account suggests that partisanship alone is insufficient to generate patterns of partisan teamsmanship and social distance, a finding generally associated with social identity theory accounts of affective polarization. Second, we posit that ideological beliefs are associated with the (mis)attribution of ideological extremity to

out-party members. Accordingly, we argue that individuals with stronger ideological commitments hold increasingly exaggerated perceptions of partisan out-groups and apply these perceptions when evaluating out-group members. Thus, our perspective indicates that ideological extremity magnifies perceptual differences across party lines (see also Lelkes 2021; Fowler 2020).

We support our argument with two analyses on the linkages between partisanship, ideology, and polarization. In the first analysis, we use an original survey experiment with a nationally representative sample of Americans to distinguish the effects of partisanship and ideology on social evaluations. We find that the attribution of ideologically extreme political views to an individual's peer significantly reduces interest in interpersonal interaction. We find limited evidence, however, that partisan group membership alone induces social polarization. In the second analysis, we build on these findings by showing that ideological extremity is strongly associated with perceptions of both partisan in-groups and out-groups. It is the most conservative Republicans and the most liberal Democrats who are most likely to misperceive the beliefs held by out-partisans. Together, our studies suggest that social polarization is largely a product of ideological disagreement rather than group conflict on the basis of partisan identity.

Partisanship and Group Evaluations

Political parties are the most salient groups for how individuals experience political phenomena (Green, Palmquist, and Schickler 2002; Iyengar, Sood, and Lelkes 2012). Citizens use partisanship as a heuristic for evaluating political candidates (Conover and Feldman 1989), forming economic assessments (Ang et al. Forthcoming), and attributing responsibility for political outcomes (Malhotra and Kuo 2008). Partisanship is also associated with nonpolitical judgments and behaviors (e.g., Iyengar and Westwood 2015; Mason 2015). Rather than reflecting retrospective calculations (Fiorina 1981), the evidence collectively suggests that partisanship separates individuals into partisan teams and influences how they perceive the world (e.g., Campbell et al. 1960; Green,

Palmquist, and Schickler 2002; Huddy, Mason, and Aarøe 2015; Mason 2015).

Recent research documents increasingly negative relationships between party membership and evaluations of out-party members. Iyengar, Sood, and Lelkes (2012) show that social distance between Democrats and Republicans has increased over the last several decades, as partisans increasingly dislike out-party members and ascribe negative traits to them. Hostile feelings toward out-party members can subsequently influence partisans' willingness to exhibit behavioral discrimination against them (Iyengar and Westwood 2015). Likewise, other scholarship finds that partisans are less likely to engage in commercial activity with (Engelhardt and Utych 2020; McConnell et al. 2018), work with (Panagopoulos et al. 2020), and date (Huber and Malhotra 2017) people who do not share their partisan identity.

Scholars have explained these phenomena by conceptualizing partisanship as an expressive social identity that generates affective reactions toward out-party members (e.g., Huddy, Mason, and Aarøe 2015; Iyengar, Sood, and Lelkes 2012; Iyengar and Westwood 2015; Mason 2015). Applying social identity theory to party membership, this perspective conceives of partisanship as “a social and psychological attachment” (Mason 2015, 129). As Huddy, Mason, and Aarøe (2015) elaborate, partisanship provides “a subjective sense of belonging to a group” in which party members seek to positively differentiate their party from the other. Because “the mere act of identifying with a political party is sufficient to trigger negative evaluations of the opposition” (Iyengar, Sood, and Lelkes 2012, 407), this perspective provides an identity-based account for affective party polarization documented in the research described above.

Distinguishing the Effect of Partisanship on Group Evaluations

Tests of identity-based accounts of affective polarization confront important inferential challenges. The primary challenge is that the nonrandom assignment of partisanship to individuals¹ makes it difficult to distinguish the effects of partisanship from other political characteristics that may also affect evaluations. For instance, membership in a political party is generally correlated with higher levels of political interest and activity (Campbell et al. 1960; Klar 2014). Moreover, Klar, Krupnikov, and Ryan (2018) demonstrate that measures of antipathy toward out-party members conflates negative partisan affect with dislike of partisan politics more generally. Partisanship is also associated with other group-based characteristics such as race, religion, and class. Evaluations of partisans could also (at least partially) reflect attitudes toward these groups (Abramowitz and Webster 2018; Ahler and Sood 2018).

In contemporary American politics, partisanship is highly correlated with issue preferences, but with two prominent exceptions most studies of affective polarization and social distance do not distinguish the effect of party membership from the effects of ideological differences. Orr and Huber (2020) address this issue with survey experiments that ask respondents to evaluate a hypothetical individual who is randomly attributed with some combination of partisan cues, social cues, and policy views. While they conclude that “a great deal of measured partisan animosity reflects disagreement about contentious issues” rather than raw partisan identity (Orr and Huber 2020, 584), the design of the experiments does not permit them to estimate the effect of partisanship and issue positions on social evaluations relative to the absence of political information.² In an

¹Fowler (2020) provides extended discussion of this point; see also Gerber, Huber, and Washington (2010) for an experimental attempt to address this issue.

²Dias and Lelkes (Forthcoming) further argue that the issues used in the Orr and Huber (2020) experiments were strongly connected with political parties, thus cueing partisanship even in the absence of an explicit partisan label.

experiment involving a hypothetical candidate, Lelkes (2021) varied the presence of partisan cues and summary ideological information before eliciting respondents' evaluation of the candidate. The results showed that evaluations were considerably more responsive to ideological information than they were to partisan cues. However, survey respondents harbor more animus toward party elites than they do toward rank-and-file party members (Druckman and Levendusky 2019), and thus it is not clear whether the results from the candidate experiment apply to evaluations of ordinary voters.

We argue that inferences about the effect of party membership on animosity toward partisan out-groups outpace the available evidence. Instead, persuasive evidence of partisan animosity should demonstrate that individuals provide systematically worse evaluations of members of partisan out-groups than they do in the absence of information about members' partisan affiliations. Our perspective is analogous to research on the role of political parties in legislative voting behavior. In this literature, scholars have argued for evaluating party effects based on how legislative outcomes would differ in the absence of political parties (see, e.g., Krehbiel 1993) and by studying the extent to which voters support co-partisans without ideological congruence on policy issues (see, e.g., Fowler 2020). While recent research makes important contributions to understanding how partisan cues and policy positions flow through one another in producing affective polarization (Dias and Lelkes Forthcoming), this scholarship does not show how party membership and ideological beliefs affect interpersonal evaluations relative to the absence of either political cue.

How Ideology Shapes Interpersonal Evaluations

We address the limitations of existing scholarship noted above and argue that perceptions of partisan outgroups are driven largely by ideological differences rather than by group membership alone. We contend that perceptions of partisan groups are rooted in perceived or actual differences in ideological beliefs. That is, individuals reach more positive assessments of their peers on

the basis of their political agreement: we feel more warmly about individuals with whom we agree than with whom we disagree. Social polarization can thus result when individuals meet others with differing political views or perceive that others have differing political views. Because partisan membership is associated with policy views (e.g., Gerber, Huber, and Washington 2010; Green, Palmquist, and Schickler 2002; Homola 2021), we expect that individuals provide more negative interpersonal assessments of individuals who belong to partisan outgroups, as prior research on partisan social identity argues (e.g., Huddy, Mason, and Aarøe 2015; Mason 2015).

We further expect that evaluations of out-party members are particularly negative for outgroup members who hold political views that are increasingly ideologically opposed to one's own. Prior work has shown that partisans perceive greater attitudinal extremity among members of the opposite party (Judd and Park 1993; Levendusky and Malhotra 2016), and other scholarship shows that higher levels of ideological extremity among the out-party are associated with affective polarization (Rogowski and Sutherland 2016; Webster and Abramowitz 2017). While information about ideological extremism may trigger negative interpersonal evaluations in the absence of explicit information about party membership, as Orr and Huber (2020) show, we argue that individuals view out-party members especially negatively when those individuals are attributed with extreme ideological beliefs. Our account posits that individuals view out-party members negatively when they hold, or are perceived to hold, views with which they disagree rather than simply on the basis of their party membership. This is consistent with evidence presented by Druckman et al. (Forthcoming), who show that most Americans view members of the other party with indifference rather than hostility, and that hostility is expressed mostly by politically engaged members of a party when asked to evaluate ideologically extreme members of the out-party.

To the extent that party membership is associated with interpersonal evaluations, we expect that these patterns are driven largely by perceptions of ideological disagreement. We argue that partisans' exaggeration of the opposing party's extreme views should be most pronounced among those who are more ideologically extreme. Ahler and Sood (2018) show that citizens tend to view

out-party members as more extreme than they are, and strong ideologues are likely to assume that others' beliefs are similarly extreme and thus project their own beliefs onto others (see, e.g., Van Boven, Judd, and Sherman 2012). In a context where conservatives have sorted into the Republican Party and liberals have sorted into the Democratic Party (Levendusky 2009; Mason 2015), therefore, ideological extremism magnifies perceptions of partisan group differences.

We report evidence from two studies to evaluate our account. The first study reports evidence to distinguish the effect of party membership and ideology on interpersonal evaluations. In contrast with expectations offered from previous studies (e.g., Dias and Lelkes Forthcoming; Huddy, Mason, and Aarøe 2015; Iyengar, Sood, and Lelkes 2012; Mason 2015), we test the hypothesis that party group membership affects interpersonal evaluations when it is accompanied by ideologically extreme beliefs but on its own has minimal effects on social polarization. The second study evaluates how an individual's ideological beliefs are associated with perceptions of partisan outgroups. Extending previous work (Ahler and Sood 2018), we test the hypothesis that more ideologically extreme partisans hold more exaggerated perceptions of members of the opposite party. Across both, evidence in favor of these hypotheses would support our argument that ideological beliefs, instead of and/or in conjunction with party membership, contribute more to social polarization than previous research has recognized.

Experimental Evidence on Social Polarization

We test our argument and distinguish the effect of partisanship and ideology on social polarization with a survey experiment. Our data come from the December 2016 wave of The American Panel Survey (TAPS), a monthly online panel survey from a national probability sample.³ Similar to the

³The sample is drawn from an address-based sampling frame and is administered online. The survey was conducted by GfK/Knowledge Networks. The December 2016 wave included 1,487 respondents. Demographic characteristics are shown in Table A.1.

design used in related studies (e.g., Orr and Huber 2020), respondents were randomly assigned to receive one of three vignettes which described the characteristics of a hypothetical new neighbor. (Complete vignette wording is shown in Table A.2.) These descriptors contained information about the person’s age and education, and indicated that the person was from a suburb outside a larger Midwestern city. We chose this information so that it would not provide explicit cues about partisan affiliation.

The *control* group (N=491) received additional non-political information about the new neighbor regarding pets and hobbies. They were also told that the neighbor frequently volunteers for local candidates’ campaigns, indicating a degree of political activism. By including this information, we ensure that comparisons between the control and treatment groups do not conflate dislike for politics or political activism in general (see, e.g., Druckman et al. Forthcoming; Klar, Krupnikov, and Ryan 2018) with dislike for specific partisan groups.

Respondents in the *partisan* treatment condition (N=497) received the same background information about the neighbor. They were also informed that the neighbor was registered as a voter for the opposing party as their own.⁴ The comparison between the control and partisan conditions allow us to compare evaluations of a neighbor who is described as affiliating with the opposite party to the counterfactual that respondents have no partisan information about the same neighbor. If out-party membership alone increases social polarization, we expect to observe more negative evaluations of the neighbor among both Republican and Democratic respondents relative to the control condition.

Respondents in the *partisanship + ideology* condition (N=499) received the same information as respondents in the *partisan* condition. They were also told that the potential neighbor subscribes to ideological viewpoints that are stereotypically associated with the neighbor’s party. These ideological perspectives were conveyed through specific policy preferences and more general

⁴We characterized respondents’ partisanship with the standard question: “Generally speaking, do you usually think of yourself as a Democrat, a Republican, an Independent, or what?”

worldviews. Democratic neighbors were characterized as believing that everyone should drive an electric car, that marijuana should be legal in all states, and that a nationalized healthcare system would improve the health of all citizens. Republican neighbors were described as believing that humans and dinosaurs walked the earth at the same time, that elementary school students should be required to recite the pledge of allegiance every morning, and that a fence should be built between the U.S. and Mexico. If individuals automatically attribute extreme issue positions to members of the out-party upon learning their partisanship, we would expect that the treatment effects of the *partisanship + ideology* condition are similar in magnitude to the treatment effects of the *partisan* condition. On the other hand, if the attribution of ideological views is the primary mechanism through which social polarization is generated, we would expect the treatment effects of the *partisan* condition to be considerably smaller in magnitude than the treatment effects of the *partisanship + ideology* condition.

Our outcome variables measure respondents' social evaluations of the hypothetical neighbor and elicit respondents' interest in ostensibly non-political interactions. After receiving the vignettes, we asked respondents whether they would be interested in befriending, hiring, dating, inviting the person over for a meal, and allowing their children to play together.⁵ Each question was measured on a five-point scale range from "strongly disagree" to "strongly agree." For simplicity, we recoded each to a binary measure that indicated whether respondents "agreed" or "strongly agreed" that they would be interested in interacting with the neighbor. Finally, all analyses use weights benchmarked to national population parameters based on the Current Population Survey (CPS) at the time of the interviews.

The design of our experiment differs in several key respects from experiments used in related scholarship. First, following guidance from Druckman and Levendusky (2019), we measure interpersonal evaluations with questions about respondents' willingness to interact with ordinary people rather than feeling thermometer ratings (Iyengar, Sood, and Lelkes 2012) which may

⁵Full question wording is in Table A.3.

capture views toward elites. Second, our primary interest is in identifying whether out-party membership is sufficient to induce social polarization, or whether affective reactions depend on the combination of out-party membership and ideological differences. Because it is *not* our goal to distinguish the effect of ideological views absent partisanship (*contra* Dias and Lelkes Forthcoming and Orr and Huber 2020), we do not include a condition that provides ideological positions without information about party membership. Third, because we *are* interested in how partisanship on its own contributes to affective polarization, we include a pure control condition while Dias and Lelkes (Forthcoming) and Orr and Huber (2020) do not. Finally, the *partisanship + ideology* condition attributes relatively extreme ideological beliefs to the neighbor. This choice contrasts with other research that has signaled ideology through more standard issues positions or more abstract ideological summaries (Dias and Lelkes Forthcoming; Lelkes 2021; Orr and Huber 2020). We believe the choice is justified given that partisans tend to exaggerate the extremism of out-party members (Ahler and Sood 2018, study 2 below). If party membership is sufficient to activate these exaggerated perceptions, we would expect similar interpersonal evaluations for respondents in the *partisan* and *partisanship + ideology* conditions. Yet if out-party membership alone does not prime respondents to attribute extreme positions to the neighbor in the vignette, as our account posits, we would expect to observe more negative evaluations from respondents in the latter condition.

Results

Figure 1 shows results for each dependent variable.⁶ The plotted points indicate the difference between the *control* condition compared with the *partisan* (shown with the triangle) and *partisanship + ideology* (shown with the circle) conditions for the outcome variables shown along the *y*-axis. The *x*-axis represents the difference in the mean share of respondents who expressed

⁶Complete summary statistics are in Table A.4.

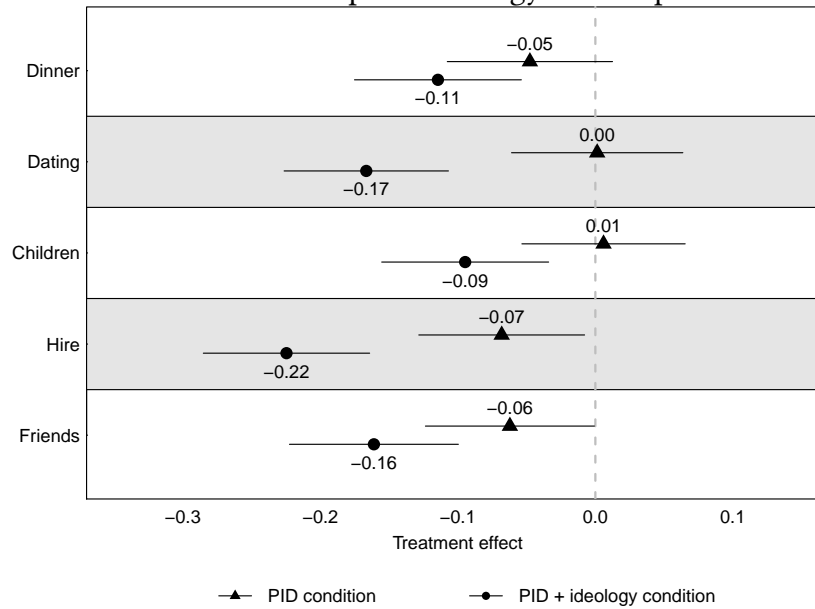
interest in interacting with the potential neighbor that, by design, is from the opposing party of the respondent. The vertical line at zero indicates the null hypothesis of no treatment effect; positive values along the x -axis indicate that larger proportions of treatment group respondents than control respondents expressed interest in interacting with the potential neighbor, while negative values indicate that smaller proportions of treatment group respondents than control respondents expressed interest in interacting with the potential neighbor.

As the figure shows, we find mixed evidence that the partisan condition significantly increased social polarization. For two of the dependent variables—hiring the neighbor as an employee and being friends with them—the effects are negative and statistically significant, indicating a six or seven percentage point decrease in social activity. The estimates for the three other dependent variables are not statistically distinguishable from zero. Notably, the effect on interest in dating—one of the most common indicators previous scholarship has used to assess social distance (e.g., Huber and Malhotra 2017)—is estimated as 0.00.

Figure 1 also shows statistically significant and substantively large effects of the *partisanship + ideology* condition on interpersonal interaction. Compared with the control group, respondents in this treatment group were substantially less interested in interacting with the hypothetical potential neighbor in all five scenarios. These differences range from -9 percentage points (allowing children to play together) to -22 percentage points (consider hiring as employee). Importantly, the effects of the *partisanship + ideology* condition are also considerably larger than and statistically distinguishable from the effects of the *partisan* condition for each outcome variable. To the extent that partisan differences contribute to social polarization, they do so mostly when paired with information about an individual's ideological worldview.

We conducted several additional analyses which we report in the Appendix. First, the patterns in Figure 1 generally apply to both Democratic and Republican respondents. The treatment effects of the *partisanship + ideology* condition were negative and statistically significant for all five dependent variables among Democratic respondents and for four of the five dependent variables

Figure 1: Effect of Partisanship and Ideology on Interpersonal Evaluations



Note: Values along the x -axis indicate the difference in proportions when comparing each of the treatment groups to the control group. The vertical line at zero indicates the null hypothesis of no treatment effect. The horizontal lines show the 95% confidence intervals.

among Republican respondents. We did find, however, that the negative treatment effects are somewhat smaller in magnitude among Republicans than Democrats. Second, the effects of the *partisan* condition were moderated somewhat by respondents' ideological extremity and partisan identity strength (see Figures A.1-A.3 and Tables A.5-A.6). Ideologically extreme respondents and respondents with stronger partisan identities reacted more negatively to the *partisan* condition than moderate respondents, for whom all the results are null.⁷

The experimental findings support our claim that differences in ideology, rather than party membership alone, are the primary driver of out-party hostility. We examine this proposed mechanism and study how respondents used the information in the vignettes to make inferences about

⁷Table A.6 and Figure A.3 demonstrate these effects are consistent when we compare those who identify as “strong” Republicans or Democrats compared to those who identify as “not strong” or “weak” partisans. For strong partisans, the effect of the *partisanship + ideology* condition was significantly in greater magnitude across all five outcomes.

the neighbor's ideological position. We asked respondents to place the neighbor on a five-point ideological scale that ranged from very liberal (1) to very conservative (5). We expect respondents in the *partisan* and the *partisanship + ideology* conditions to provide more ideologically extreme assessments of the neighbor relative to respondents in the control condition. For Republican respondents, this would manifest in a more liberal rating of the neighbor, and for Democratic respondents this would manifest in a more conservative rating of the neighbor. Given our account, we further expect that respondents in the *partisanship + ideology* condition would evaluate their neighbor as more extreme than either the control or *partisan* conditions.

Figure 2 shows the results. Negative numbers along the x -axis indicate that respondents provided more liberal evaluations relative to the control group and positive numbers indicate that respondents provided more conservative evaluations relative to the control group. The vertical line at zero again indicates the null hypothesis of no difference in evaluations relative to the control group. The results for Republican respondents (evaluating Democratic neighbors) are shown with the dark circles and the results for Democratic respondents (evaluating Republican neighbors) are shown with dark triangles.

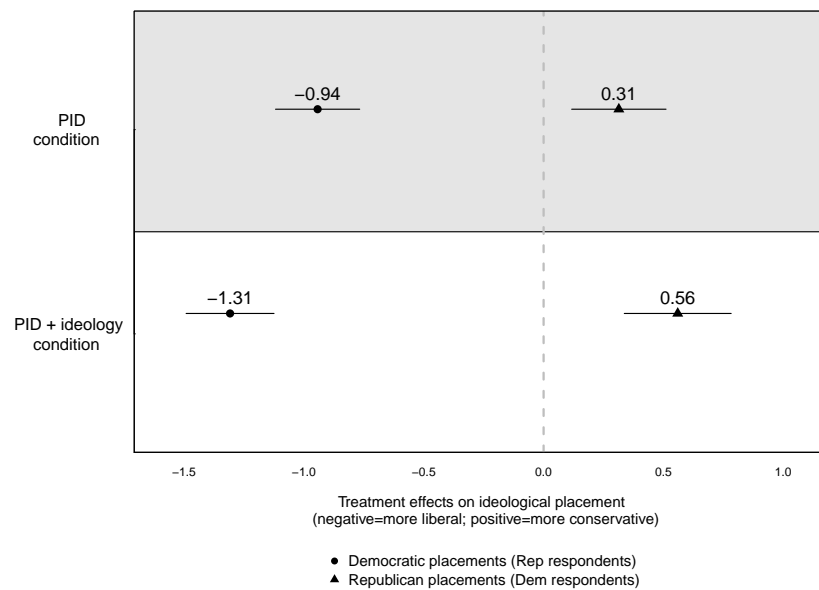
Republican and Democratic respondents in the control condition provided generally similar ideological evaluations of the potential neighbor. On the five-point scale, Democrats placed the neighbor at 3.1 and Republicans placed the neighbor at 3.4. Absent specific partisan or ideological information about the neighbor, respondents therefore placed the individual referenced in the vignettes around the center of the ideological scale.

The top panel shows the results for the *partisan* condition. Overall, Republican respondents placed the Democratic neighbor almost a full point in the more liberal direction. The results for placements by Democratic respondents were about one-third the magnitude, indicating that the *partisan* condition increased their evaluation of the neighbor's conservatism by 0.31 points on the five-point scale. Absent any policy or other ideological information, therefore, we find that partisan affiliation polarizes how partisan respondents evaluated the ideological orientation of the

neighbor referenced in our vignettes.

The bottom panel shows the results for the *partisanship + ideology* condition. For both Republican and Democratic respondents, we find that this treatment significantly increased the perceived extremity of the potential neighbor. And for both parties, the magnitude of the effect is roughly 50 percent larger relative to the effect of the partisan condition. Among both Republican and Democratic respondents, therefore, partisanship, particularly when coupled with ideological information, significantly affected their evaluation of the extremism of the potential neighbor's ideology.

Figure 2: Effect of Partisanship and Ideology on Evaluations of Ideology



Note: Values along the x -axis indicate the treatment effects of the conditions shown on the y -axis relative to the control group. Negative numbers indicate more liberal placements and positive numbers indicate more conservative placements. The horizontal lines show the 95% confidence intervals.

Our experimental findings provide new evidence about how partisanship and ideology affect interpersonal evaluations. In contrast to claims about the effect of party membership on affective polarization, partisanship has relatively small, and mostly statistically insignificant, effects on social polarization. In other words, absent the presence of ideological cues, partisanship itself has a relatively limited effect on social interaction (contra Orr and Huber 2020). Consistent with

our argument, the experimental results indicate that the invocation of ideological characteristics significantly increases social distance, at least when these views are paired with information about party membership. Overall, these findings suggest that previous research might have overstated the contributions of partisanship to social polarization and indicate that differences in ideological beliefs are more substantial contributors to interpersonal evaluations.

How Partisans View the Other

The experiment in the preceding section shows how the invocation of issue positions causes individuals to create social distance between themselves and out-party members. In a second analysis, we study the *sources* of perceptions of partisan outgroups. In particular, we focus on how an individual's own ideological beliefs are associated with perceptions of ideological extremism among supporters of the opposing party. To the extent some individuals hold systematically more extreme perceptions of members of the opposing party, the predictors of these perceptions rather than partisanship per se may help explain partisan affective polarization documented in previous survey research.

We used data from the March 2014 and May 2014 waves on TAPS, on which we fielded a battery of questions to measure perceptions of partisan groups. Our sample includes 1,301 panelists who completed both waves. This subset of panelists, which forms our sample for this analysis, is demographically representative of the U.S. population. Summary statistics are shown in Tables B.1 and B.2.

We asked ten questions to measure partisan beliefs in the March 2014 wave. For each party, our questions were intended to tap into perceptions of lifestyles and cultural values in addition to policy attitudes. Five statements corresponded to perspectives that may be typically associated with Republicans and included topics such as creationism, gun ownership, and nationalism (e.g., "We should build a fence between the United States and Mexico."). The other five statements aimed

to capture perspectives more likely to be associated with Democrats including topics such as paternalism, tax increases, and eco-friendliness (e.g., “This country would be better if every citizen drove an electric car.”). The complete question wordings are in Table B.3.

Though these statements are not exhaustive, they were inspired by messages from political actors that attributed specific qualities to their opponents and media descriptions of prominent partisans. For example, Donald Trump made the construction of a wall between the U.S. and Mexico a prominent component of his 2016 presidential campaign. Similarly, while president, Barack Obama not only made it his goal to put one million all-electric cars on the road by 2015 but also promised to buy one himself after his presidency.⁸ We used responses to these statements, which we term *first-order beliefs*, to calculate the proportions of Democrats, Republicans and Independents that agreed with each.

Table 1 presents respondents’ first-order beliefs and shows the percentages of Democrats, Independents, and Republicans that agreed with each item. The top panel shows rates of agreement with the items intended to assess typical perceptions of Republicans and the bottom panel shows rates of agreement with the items expected to elicit commonly-held Democratic perceptions. Each item generally performed as expected, with greater rates of agreement with the Republican items among self-identified Republicans. For example, 42.5% of all Republican respondents agreed with the statement that “this country would be safer if every law-abiding citizen possessed a firearm,” while agreement was much lower among Independents (25.1%) and Democrats (14.2%). Similarly, for the five Democratic items we observe the highest levels of agreement among self-identified Democrats. For example, 38.7% of Democrats agreed that “this country would be better if every citizen drove an electric car,” while smaller percentages of Independents (22%) and Republicans (11.4%) agreed. Furthermore, for nine of the 10 questions Independents agreed with each statement at rates in between those for Democrats and Republicans.

Columns 4 and 5 of Table 1 present results of statistical tests of differences in agreement

⁸See <http://wapo.st/2xAIY4I>, <http://bit.ly/2b8GgVY>, and <http://bit.ly/1Uisr5d>.

between Democrats and Republicans. Column 4 presents the absolute difference in percentage points, and Column 5 reports the accompanying *t*-statistics for differences in means tests. The differences in agreement between Democrats and Republicans are statistically significant for nine of the ten items. The only statement that does not exhibit significant differences is: “The federal government should impose a ban on a sale of soda.”⁹ Moreover, for all of our items, the differences are of substantively important magnitudes. Excluding the soda ban item, the differences range from 9.3 percentage points on the question regarding humans and dinosaurs to a striking 47 percentage points when asked about nationalized health care. Given the variation in rates of agreement among partisans and the differences between parties, the first-order beliefs reported in Table 1 provide good leverage for studying perceptions of partisan groups.

Table 1: Descriptive Statistics: First-Order Beliefs

	Democrat Agreement	Independent Agreement	Republican Agreement	$ D - R $	t-statistic
All have guns	14.2	25.1	42.5	28.3	6.19
Humans and dinosaurs together	19.1	18.5	28.4	9.3	2.17
Homosexuality threatens	13.2	17.5	34.5	21.3	5.04
Require pledge	49.1	52.9	72.6	23.5	4.59
Border fence	27.3	40.9	53.3	26.0	5.44
Electric car	38.7	22.0	11.4	27.3	-7.19
Legal marijuana	43.3	41.4	24.7	18.6	-4.19
Ban on soda	12.1	7.0	6.9	5.2	-1.49
All pay more taxes	19.3	12.0	6.8	12.5	-3.84
Nationalized health care	63.2	33.8	16.2	47.0	-10.24

Note: The agreement columns report the percentage of Democrats, Independents, and Republicans who indicated agreement with the respective item. The second to last column reports the absolute difference in mean level of agreement between Democrats and Republicans. Bold differences indicate statistical significance at the 95% confidence level.

The May 2014 wave asked panelists to indicate the percentages of Democrats/Republicans they perceived to agree with each of the statements introduced in the March wave. Methodologically, the two-month lag between asking respondents about their own beliefs and their perceptions of partisan groups’ beliefs helps to minimize potential biases from anchoring or reference effects.

⁹Since agreement with this item is relatively low among Democrats, Republicans and Independents, it is not surprising that the existing difference is not significant.

Respondents represented their perceptions of partisan groups' beliefs by choosing from five response categories that corresponded to their perceived rates of agreement with each statement: 0-20 percent, 21-40 percent, 41-60 percent, 61-80 percent, and 81-100 percent. We coded these responses on a five-point scale, where larger numbers represent a respondent's perception that greater proportions of the partisan group share that belief or attitude. For example, when Democratic panelists were asked to indicate what percentage of Republicans they believe agreed with the statement "Elementary students should be required to recite the pledge of allegiance every day," a larger number indicates that they perceived substantial agreement with that statement among Republicans. This measure of *second-order beliefs* characterizes how strongly respondents attributed these perceptions to partisan groups.¹⁰ These second-order beliefs are our primary dependent variable.

A potential limitation is that biased perceptions of the outgroup could be caused by a desire to express ideological extremity of the opposite party. Additionally, our strategy cannot directly address the causes of variation in the independent variable, and a key threat to inference is whether there is some third factor that causes both own (apparent) extremity and (apparent) views about the extremity of the other party. Yet there are reasons to believe the survey responses are sincere and not endogenous (see also Berinsky 2018). By asking about multiple extreme beliefs, we observe a tremendous amount of variation as, for example, ideologically extreme respondents report holding some, but not all, of the partisan misperceptions. This variation suggests sincerity on the part of the respondents; homogeneity in extremity would suggest more affective (perhaps signaling) behavior.

¹⁰Our use of the term *second-order beliefs* is consistent with its use in psychology to explain "what people think about other people's thoughts" (Perner and Wimmer 1985), but is distinct from its use in game theoretic scenarios where a strategic actor develops beliefs about the strategy another actor expects her to use.

Ideology and Partisan Perceptions

Do Republicans [Democrats] overestimate the prevalence of Democrats [Republicans] who hold extreme beliefs? Our data indicates that outpartisans overestimate the beliefs of their counterparts with high frequency. For all but one item (the legalization of marijuana), the modal Republican overestimated the true percentage of Democrats who agree with the issue. This overestimation is most pronounced when considering the attitudes about taxes. Only 19% of Democrats agreed with raising everyone’s taxes, while more than 80% of Republicans believed the percentage to be higher than that figure. In fact, the modal Republican responded that over 80% of Democrats agreed with the statement. Democrats followed a similar pattern: the modal Democrat overestimated the “true” percentage for four of the five items.¹¹ Thus, we see that both Republicans and Democrats overestimate the true percentage of their counterpartisans who hold an extreme belief.¹²

Partisan (Mis)Perceptions and Ideological Extremity

We now test our prediction that an individual’s own ideological beliefs are associated with increased exaggerations in perceptions of the opposite party. Specifically, we expect that those who are themselves ideologically extreme will be most likely to exaggerate the extremity of out-party supporters. We characterize ideological extremity with a measure of ideology based on respondents’ agreement with a series of 13 policy statements known to load highly on a single

¹¹Even for the one issue in which the modal Democrat matched with the correct percentage (Dinosaurs and humans coexisting), nearly 40% of Democrats still overestimated the true percentage.

¹²Figure C.1 and its discussion in the Appendix provide more information on the descriptive statistics of perceptions of counterpartisans’ beliefs.

ideological dimension (Claassen, Tucker, and Smith 2015).¹³ The scores have a mean of zero (SD=1); because higher values of this measure represent more conservative preferences and lower values indicate liberal preferences, we refer to this measure as *Conservatism*.

We regress each respondent's second-order beliefs of out-party members on their ideological leanings using ordinary least squares. For respondents identifying as partisans, we focus on the five items that relate to the *opposite* party. The unit of analysis is respondent *i*'s reported perception of out-party views on question *j*. For example, for the statement "this country would be safer if every law-abiding citizen possessed a firearm", $Perception_{ij}$ is included for every Democrat in our sample and takes on values between 1 – 5, where higher values indicate they believe a larger share of Republicans agreed with the statement. Our model is as follows:

$$Perception_{ij} = \beta_1 + \beta_2 Conservatism_i + \gamma \mathbf{X}_i + \delta_j + \varepsilon_{ij},$$

where $Conservatism_i$ is our measure of operational ideology for respondent *i*.¹⁴ To address potential question-specific variation in perceptions, we include indicator variables for each item, represented by δ_j . Additionally, $\gamma \mathbf{X}_i$ captures the effect of a series of control variables, including political knowledge and political interest.¹⁵ Respondents with high political knowledge and

¹³Factor loadings are in Table B.4 and question wordings are in Table B.5. Figure B.1 displays the ideology scores by partisan identification.

¹⁴Our measure of ideology could describe individuals' sincere policy beliefs or their ideological constraint (Broockman 2016). Political knowledge might also influence responses to our policy questions and evaluations of members of the out-party (Bullock 2011). Our results are nearly identical to those reported in Table 2 when studying these relationships among low-knowledge respondents, and we find no evidence that knowledge systematically moderates the relationship between reported policy views and out-party perceptions.

¹⁵Political knowledge is measured as the number of correct answers in a ten-item battery covering

interest may hold exaggerated second-order beliefs about members of the out-party because they are more familiar with the opinions and beliefs of out-party politicians and supporters due to their exposure to elite rhetoric. We also include a broad set of socio-demographic controls.¹⁶ We cluster standard errors by respondent.

The results are presented in Table 2, where Column 1 shows Republicans' views of Democrats and Column 2 explains Democrats' views of Republicans. Consistent with our expectations, the coefficient estimate for *Conservatism* is statistically significant in both models. The positive coefficient estimate in Column 1 indicates that more conservative Republicans perceived higher rates of agreement among Democrats with the Democratic items, while the negative coefficient estimate in Column 2 suggests that more liberal Democrats perceived that larger percentages of Republicans subscribed to the Republican items.

The third column of Table 2 includes both Democrats' and Republicans' views of the out-party. Because *Conservatism* is standardized at 0, we use its absolute value as a measure of *Ideological Extremity*. High values of this variable indicate respondents that are very liberal or very conservative, while lower values indicate respondents with more ideologically moderate views. The positive and significant coefficient estimate shows that more ideologically extreme partisans perceived higher levels of agreement with items among their respective out-party.¹⁷

political affairs and American government. Political interest is measured with the question "In general, how interested are you in politics and public affairs?" with responses on a four-point scale ranging from "Not at all interested" to "Very interested."

¹⁶Results for the full set of control variables can be found in Table B.6.

¹⁷We also examined panelists' perceived level of agreement with these statements among *in-party* members. Here, *Conservatism* is positively associated with the outcome variable for Republicans and negatively so for Democrats, indicating that more ideologically extreme partisans believe greater proportions of their in-party members subscribe to the statements. While these estimates are statistically significant, they are weaker in magnitude than those in the main analysis. The

Table 2: Partisans' Views of Out-Party Members

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	0.336* (0.073)	-0.370* (0.069)	
Ideological Extremity			0.459* (0.066)
Political Knowledge	0.038 (0.029)	0.060* (0.022)	0.047* (0.018)
Political Interest	0.275* (0.077)	0.097 (0.056)	0.168* (0.045)
Constant	2.224* (0.508)	2.530* (0.311)	2.005* (0.268)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1513	2480	3993
<i>N</i> (Respondents)	323	537	860
<i>R</i> ²	0.23	0.19	0.20

Note: Table entries are linear regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is Perception_{ij} , the degree to which a respondent i believes members of the out-party agreed with item j as described in the text. Additional socio-demographic controls (income, education, sex, race, age, Metropolitan Statistical Area, and US Census region) and question indicators are included but not reported. * $p < 0.05$

The estimated coefficients for the control variables are also of substantive interest. More politically interested and sophisticated respondents tended to believe that a larger share of members of the other party agreed with that party's items.¹⁸ Though the exact level of statistical significance differences between the coefficients are statistically significant at $p < .10$ for columns 1 and 2 and at $p < .05$ for column 3 (likely due to the increased statistical power from pooling all respondents), indicating that ideological extremity is a stronger predictor of out-party perceptions than in-party perceptions. The full results are shown in Table B.7 in the Appendix.

¹⁸We also estimated models treating political interest as a series of dummy indicators. The estimated coefficients for the dummy variables indicated a near-linear association with the

varies across these estimates, the results generally suggest that greater attentiveness to politics and public affairs among partisans is associated with attributing these beliefs to members of the out-party at greater rates.¹⁹

Our results are robust to a number of different model specifications and measurement choices. We estimated models that included an indicator for whether a respondent is a “strong partisan” to explore whether our results are driven by strength of party identification rather than ideological extremity (Table B.9). We also substituted a traditional seven-point symbolic ideology variable for our measure of ideology (Tables B.10 and B.11). To disentangle ideological extremity from one’s ideological identity, we controlled for both symbolic and operational ideology in the same models (Tables B.12 and B.13). We estimated models in which we accounted for the partisan’s own position on each statement to explore the possibility that respondents may hold exaggerated perceptions in ways that varied systematically with their position on that statement (Table B.14). We also measured party identification in which *leaners* (people thinking of themselves as closer to one of the parties in a follow-up question) are classified as partisans rather than as Independents (Table B.15). Finally, we created an indicator for responses whose partisan identities were aligned with their ideologies to explore whether our results are driven by partisan-ideological sorting (Table B.16).²⁰ We also estimated models that included the different controls for strong partisanship,

dependent variable.

¹⁹We also conducted this analysis on Independents. Consistent with our argument, more liberal Independents tended to believe that greater proportions of Republicans agreed with the Republican items we asked, whereas more conservative Independents believed that more Democrats agreed with the Democratic items. This is consistent with recent research that finds that citizens who may be embarrassed about their partisan attachments identify as Independents despite holding beliefs similar to partisans (Klar and Krupnikov 2016). Results may be found in Appendix Table B.8

²⁰We created an indicator, *Sorted*, for partisans who fell on their own ideological side of the

symbolic ideology, and partisan ideological sorting all at the same time (Table B.17). These models produce consistently strong evidence that ideological extremity is associated with exaggerated perceptions of partisan groups.

In addition, because our outcome measure is a five-category variable, we replicated all analyses using ordered logit models (Table B.18).²¹ And, while our analyses pool all perception items together, we estimated a series of models that focused on one individual statement at a time (Tables B.19 and B.20). Moreover, in December 2016 we also asked respondents to indicate their answers on a 100-point scale (Tables B.21 and B.22). These robustness checks also support the findings above.

Figure 3 illustrates the substantive relationship between ideology and partisan perceptions and shows predicted levels of our outcome variable $Perception_{ij}$ across the range of values of operational ideology. Based on Models 1 and 2 in Table 2, Figure 3 displays the differences in perceptions of members of the out-party.²² The y-axis corresponds to the categorical outcome measure, where a value of ‘2’ indicates a belief that between 21 and 40 percent of out-partisans agree with a given item and ‘4’ indicates a belief that between 61 and 80 percent of out-partisans agree with a given item. The dashed line is the predicted perception among Democrats toward Republicans and the solid line is the predicted perception among Republicans toward Democrats. The shaded regions represent the 95% confidence intervals. Figure 3 suggests that the most

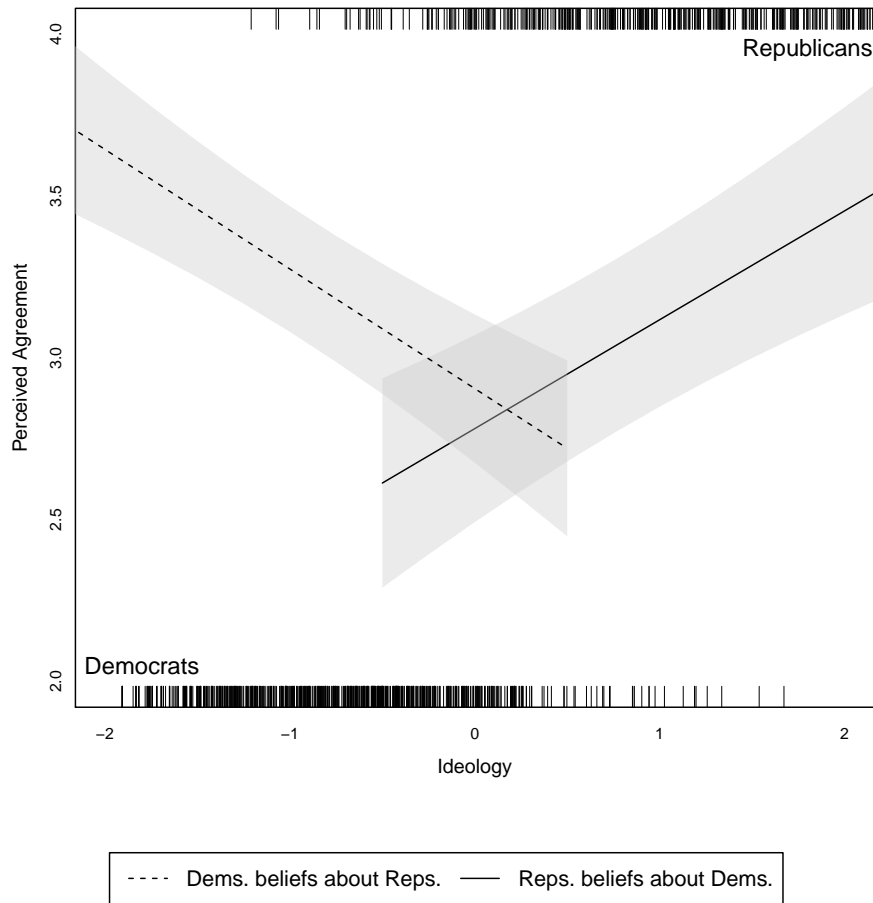
Conservatism measure. If a Republican (Democrat) scored on the positive (negative) side of 0 in our standardized operational conservatism measure, she was coded as 1, or “sorted.” Likewise, if a Republican [Democrat] scored on the negative (positive) side of 0 in our conservatism variable, she was coded as 0, or “unsorted.”

²¹Since the findings of these ordered logit models were consistent with a linear modeling strategy, we present the results from the linear models for a more straightforward interpretation.

²²For the purpose of these predictions, all other continuous covariates were held at their mean scores, while categorical and dummy variables were held at their modal values.

conservative Republicans and the most liberal Democrats perceive that members of the opposite party agree with a respective item a full category higher than their more moderate co-partisans, all else equal. That is, given two Republicans who are identical in all respects, with the exception of one being a far-right conservative and the other being an ideological moderate, our model predicts that the more conservative individual will believe about 20 percent more Democrats agree with a given Democratic item.

Figure 3: Predicted Second-Order Beliefs for Partisans



Note: Values along the x -axis indicate respondents' operational ideology, values along the y -axis the predicted values of second-order beliefs. The shaded areas show 95% confidence intervals.

We also examined whether respondents identified the *correct* proportion of out-party members who agreed with each survey item. These results are shown in the Appendix in Table C.1. For each item, we created an indicator for whether a respondent correctly identified the percentage

category that contained the observed level of agreement with the partisan item.²³ We find that the likelihood of identifying the correct proportion of agreement among partisan out-groups is significantly related to the respondent's own ideology, where more ideologically extreme views are associated with less accurate perceptions of the other party.

The results from this analysis support our argument that perceptions of partisan out-groups are substantially influenced by ideology. Among both Democrats and Republicans, ideological extremity is associated with greater inaccuracy and increased exaggeration in perceptions of the out-group. Consistent with our theoretical perspective and the experimental results discussed above, we find that ideology plays an important role in explaining exaggerated perceptions.

Conclusion

Several generations ago, responsible party theorists emphasized the centrality of political parties to modern democracy (e.g., Schattschneider 1942). Today, however, membership in and identification with political parties is often attributed with a host of social ills, including support for undemocratic behavior (Graham and Svobik 2020) and increased antipathy toward members of the opposite party (e.g., Dias and Lelkes Forthcoming; Huddy, Mason, and Aarøe 2015; Iyengar, Sood, and Lelkes 2012; Iyengar and Westwood 2015; Mason 2015). On this latter score, our results suggest that previous research has overstated the consequences of partisanship for affective polarization. Instead, partisans are most likely to harbor negative views of out-party members when those individuals hold, or are believed to hold, ideological views that are out of step with their own. In contrast with research that argues affective polarization is caused primarily by partisan identity

²³For example, 28.4% of all Republicans agreed with the statement that “Humans and dinosaurs walked the earth at the same time.” A Democratic respondent who is asked about the percentage of Republicans they believed agreed with that statement are coded as 1 if they responded that 21-40% of Republicans agreed with this statement, and 0 otherwise.

(Dias and Lelkes Forthcoming), our findings indicate that social distance is generated largely by individuals' ideological commitments. Our analyses suggest that accounting for policy views in the contemporary electorate will allow researchers to more fully understand the dynamics that characterize mass politics.

Our findings contribute new insights about the potential consequences of contemporary partisan polarization for mass political behavior. In particular, our results suggest that increased partisan polarization may be associated with the use of increasingly exaggerated partisan caricatures. As the public forms perceptions of partisan groups at increased rates and with decreased accuracy, the “pictures in our heads” may in fact make it more difficult for citizens to relate to one another across the partisan aisle. These patterns of results suggest a potential explanation for why increased polarization at the elite level has led to increased social polarization among partisans in the mass public. They also suggest an important link between increased ideological extremity and decreased affect toward political out-groups.

By design, our research has some important limitations. First, our survey experiment asked respondents to evaluate the potential for interpersonal interaction based on the information they received in a hypothetical setting. We are less certain how the experimental results generalize into real-world settings in which, for instance, social norms could serve as countervailing influences on the application of partisan perceptions. Second, we considered a small number of items that may be salient for forming perceptions of partisans. Third, our TAPS data mostly represent a single snapshot in time and limit our ability to make stronger conclusions about the causal relationship between partisan polarization and perceptions. It is unclear whether the nature of partisan perceptions has intensified in the contemporary era relative to a generation or two ago. Fourth, the design of our studies does not permit us to conduct convincing mediation analyses to evaluate the mechanisms implicated by our theory. More systematic research, both experimental and observational, is necessary to more definitively answer these questions.

At the same time, our research raises several important questions about the nature of partisan

perceptions and their implications. Perhaps most importantly, future research should interrogate the possibility of correcting exaggerated (mis)perceptions about partisans. For instance, Ahler (2014) shows that correcting exaggerated perceptions of ideological polarization reduced respondents' self-reported levels of ideological extremism. To the degree that partisan perceptions inhibit social interaction across partisan lines, efforts to increase the accuracy of partisan perceptions may therefore prove fruitful for achieving greater levels of empathy, reduced affective polarization, and more consensual political outcomes. Finally, our research does not address temporal dynamics in the formation and deployment of partisan stereotypes, though we suspect that they emerge, evolve, and are displaced over time as the parties themselves change (e.g., Green, Palmquist, and Schickler 2002). These questions are all important for future research.

References

- Abramowitz, Alan I., and Steven W. Webster. 2016. "The Rise of Negative Partisanship and the Nationalization of U.S. Elections in the 21st Century." *Electoral Studies* 41: 12–22.
- Abramowitz, Alan I., and Steven W. Webster. 2018. "Negative Partisanship: Why Americans Dislike Parties But Behave Like Rabid Partisans." *Advances in Political Psychology* 39: 119–135.
- Ahler, Douglas J. 2014. "Self-Fulfilling Misperceptions of Public Polarization." *Journal of Politics* 76: 607–620.
- Ahler, Douglas J., and Gaurav Sood. 2018. "The Parties in Our Heads: Misperceptions About Party Composition and Their Consequences." *The Journal of Politics* 80(3): 964–981.
- Alduncin, Alex, David C. W. Parker, and Sean M. Theriault. 2017. "Leaving on a Jet Plane: Polarization, Foreign Travel, and Comity in Congress." *Congress & the Presidency* 44: 179–200.
- Ang, Zoe, Andrew Reeves, Jon C. Rogowski, and Arjun Vishwanath. Forthcoming. "Partisanship, Economic Assessments, and Presidential Accountability." *American Journal of Political Science* .
- Berinsky, Adam J. 2018. "Telling the Truth about Believing the Lies? Evidence for the Limited Prevalence of Expressive Survey Responding." *Journal of Politics* 80(1): 211–224.

- Broockman, David E. 2016. "Approaches to Studying Policy Representation." *Legislative Studies Quarterly* 41: 181–215.
- Bullock, John G. 2011. "Elite Influence on Public Opinion in an Informed Electorate." *American Political Science Review* 105: 496–515.
- Campbell, Angus, Philip E. Converse, Warren E. Miller, and Donald E. Stokes. 1960. *The American Voter*. New York: Wiley.
- Claassen, Christopher, Patrick Tucker, and Steven S. Smith. 2015. "Ideological Labels in America." *Political Behavior* 37: 253–278.
- Conover, Pamela J., and Stanley Feldman. 1989. "Candidate Perception in an Ambiguous World: Campaigns, Cues, and Inference Processes." *American Political Science Review* 33(4): 912–940.
- Desmarais, Bruce A., Vincent G. Moscardelli, Brian F. Schaffner, and Michael S. Kowal. 2015. "Measuring Legislative Collaboration: The Senate Press Events Network." *Social Networks* 40: 43–54.
- Dias, Nicholas, and Yphtach Lelkes. Forthcoming. "The Nature of Affective Polarization: Disentangling Policy Disagreement from Partisan Identity." *American Journal of Political Science* .
- Druckman, James N., and Matthew S. Levendusky. 2019. "What Do We Measure When We Measure Affective Polarization?" *Public Opinion Quarterly* 83(1): 114–122.
- Druckman, James N., Samara Klar, Yanna Krupnikov, Matthew Levendusky, and John Barry Ryan. Forthcoming. "(Mis-)Estimating Affective Polarization." *Journal of Politics* .
- Engelhardt, Andrew M., and Stephen M. Utych. 2020. "Grand Old (Tailgate) Party? Partisan Discrimination in Apolitical Settings." *Political Behavior* 42: 769–789.
- Fiorina, Morris P. 1981. *Retrospective Voting in American National Elections*. New Haven: Yale University Press.
- Fowler, Anthony. 2020. "Partisan Intoxication or Policy Voting?" *Quarterly Journal of Political Science* 15(2): 141–179.
- Gentzkow, Matthew, Jesse M. Shapiro, and Matt Taddy. 2016. "Measuring Polarization in High-Dimensional Data: Method and Application to Congressional Speech." *NBER Working Paper* 22423.

- Gerber, Alan S., Gregory A. Huber, and Ebonya Washington. 2010. "Party Affiliation, Partisanship, and Political Beliefs: A Field Experiment." *American Political Science Review* 104: 720–744.
- Graham, Matthew H., and Milan W. Svobik. 2020. "Democracy in America? Partisanship, Polarization, and the Robustness of Support for Democracy in the United States." *American Political Science Review* 114(2): 392–409.
- Green, Donald, Bradley Palmquist, and Eric Schickler. 2002. *Partisan Hearts and Minds*. New Haven, Conn.: Yale University Press.
- Homola, Jonathan. 2021. "Partisanship and perceived threats about immigration." *Party Politics* 27(5): 977–982.
- Huber, Gregory A., and Neil Malhotra. 2017. "Political Homophily in Social Relationships: Evidence from Online Dating Behavior." *The Journal of Politics* 79(1): 269–283.
- Huddy, Leonie, Lilliana Mason, and Lene Aarøe. 2015. "Expressive Partisanship: Campaign Involvement, Political Emotion, and Partisan Identity." *American Political Science Review* 109: 1–17.
- Iyengar, Shanto, and Sean J. Westwood. 2015. "Fear and Loathing across Party Lines: New Evidence on Group Polarization." *American Journal of Political Science* 59: 690–707.
- Iyengar, Shanto, Gaurav Sood, and Yphtach Lelkes. 2012. "Affect, Not Ideology: A Social Identity Perspective on Polarization." *Public Opinion Quarterly* 76: 405–431.
- Iyengar, Shanto, Yphtach Lelkes, Matthew Levendusky, Neil Malhotra, and Sean J. Westwood. 2019. "The Origins and Consequences of Affective Polarization in the United States." *Annual Review of Political Science* 22: 129–146.
- Judd, Charles M., and Bernadette Park. 1993. "Definition and Assessment of Accuracy in Social Stereotypes." *Psychological Review* 100: 109–128.
- Klar, Samara. 2014. "Identity and Engagement among Political Independents in America." *Political Psychology* 35(4): 577–591.
- Klar, Samara, and Yanna Krupnikov. 2016. *Independent Politics: How American Disdain for Parties Leads to Political Inaction*. New York: Cambridge University Press.

- Klar, Samara, Yanna Krupnikov, and John B. Ryan. 2018. "Affective Polarization or Partisan Disdain?: Untangling a Dislike for the Opposing Party from a Dislikes of Partisanship." *Public Opinion Quarterly* 82(2): 379–390.
- Krehbiel, Keith. 1993. "Where's the Party?" *British Journal of Political Science* 23: 235–266.
- Lelkes, Yphtach. 2021. "Policy Over Party: Comparing the Effects of Candidate Ideology and Party on Affective Polarization." *Political Science Research and Methods* 9: 189–196.
- Levendusky, Matthew. 2009. *The Partisan Sort: How Liberals Became Democrats and Conservatives Became Republicans*. Chicago: University of Chicago Press.
- Levendusky, Matthew S., and Neil Malhotra. 2016. "(Mis)perceptions of Partisan Polarization in the American Public." *Public Opinion Quarterly* 80: 378–391.
- Malhotra, Neil, and Alexander G. Kuo. 2008. "Attributing Blame: The Public's Response to Hurricane Katrina." *Journal of Politics* 70: 120–135.
- Mason, Lilliana. 2015. "I Disrespectfully Agree': The Differential Effects of Partisan Sorting on Social and Issue Polarization." *American Journal of Political Science* 59: 128–145.
- McConnell, Christopher, Yotam Margalit, Neil Malhotra, and Matthew Levendusky. 2018. "The Economic Consequences of Partisanship in a Polarized Era." *American Journal of Political Science* 62(1): 5–18.
- Orr, Lilla V., and Gregory A. Huber. 2020. "The Policy Basis of Measured Partisan Animosity in the United States." *American Journal of Political Science* 64(3): 569–586.
- Panagopoulos, Costas, Donald P. Green, Jonathan Krasno, Michael Schwam-Baird, and Kyle Endres. 2020. "Partisan Consumerism: Experimental Tests of Consumer Reactions to Corporate Political Activity." *The Journal of Politics* 82(3): 996–1007.
- Perner, Josef, and Heinz Wimmer. 1985. "'John Thinks That Mary Thinks That . . .': Attribution of Second-Order Beliefs by 5- to 10-Year-Old Children." *Journal of Experimental Child Psychology* 39: 437–471.
- Rogowski, Jon C., and Joseph L. Sutherland. 2016. "How Ideology Fuels Affective Polarization." *Political Behavior* 38: 485–508.
- Schattschneider, E.E. 1942. *Party Government*. New York: Rinehart & Company.

- Simas, Elizabeth N., Scott Clifford, and Justin H. Kirkland. 2020. "How Empathic Concern Fuels Political Polarization." *American Political Science Review* 114(1): 258–269.
- Van Boven, Leaf, Charles M. Judd, and David K. Sherman. 2012. "Political Polarization Projection: Social Projection of Partisan Attitude Extremity and Attitudinal Processes." *Journal of Personality and Social Psychology* 103: 84–100.
- Webster, Steven W., and Alan I. Abramowitz. 2017. "The Ideological Foundations of Affective Polarization in the U.S. Electorate." *American Politics Research* 45(4): 621–647.

Appendix

“Through the Ideology of the Beholder: How Ideology Shapes Perceptions of Partisan Groups”

This Appendix provides additional analysis and figures for “Through the Ideology of the Beholder: How Ideology Shapes Perceptions of Partisan Groups” and includes 3 main parts:

- **Part A** TAPS Study – Experiment (p. 2-11) : this part contains 3 figures and 6 tables that provide additional information regarding the paper’s experimental analysis that examines the consequences of partisan (mis)perceptions.
- **Part B** TAPS Study – Study 2, Perceptions (p. 12-35) : this part contains 1 figure and 22 tables that provide additional information regarding the paper’s analysis of partisan (mis)perceptions and ideological extremity.
- **Part C** TAPS Study – Correct Assessment (p. 37-42) : this part contains 2 tables and 1 figure that provide additional information regarding an additional set of analyses that examines whether respondents identified the correct proportion of out-party members who agreed with each survey item.

A TAPS Study – Experiment

The following tables and figures provide additional background information regarding our subject pool, the measurement and distributions of key variables, and also the results from a series of robustness checks. More specifically, Table A.1 provides descriptive statistics of the socio-demographic and political variables that define our sample. Table A.2 contains the full vignette wording, while Table A.3 presents the full question wording for the outcome variables.

Table A.4 provides summary statistics for the outcome variables across the different experimental treatment groups. Figures A.1, A.2, and A.3 show the treatment effects by respondents' ideological extremity (A.1), partisan identification (A.2), and strength of partisan identity (A.3) respectively. Finally, Tables A.5 and A.6 present heterogeneous effects by partisan identification and strength of partisanship.

Table A.1: Descriptive Statistics (TAPS Experiment)

Variable	N	Mean	Standard Deviation	Minimum	Maximum
Female	1487	0.51	0.50	0.00	1.00
White	1487	0.78	0.42	0.00	1.00
Age	1460	56.46	15.56	18.00	93.00
Education	1477	11.30	1.79	3.00	15.00
Income	1400	7.00	3.66	1.00	16.00
Ideology	1480	4.17	1.68	1.00	7.00
Democrat	1487	0.36	0.48	0.00	1.00
Republican	1487	0.31	0.46	0.00	1.00

Table A.2: Vignette Question Wording (TAPS Experiment)

Condition	Vignette
Control	<p>Suppose a new person moves into your neighborhood. The person is approximately 45 years old, has a college degree, and grew up in a suburb outside of a large Midwestern city.</p> <p>This person has a dog, enjoys being physically active, follows college sports and frequently volunteers to work in local candidates' campaigns.</p>
Partisan (R)	<p>Suppose a new person moves into your neighborhood. The person is a registered Republican, approximately 45 years old, has a college degree, and grew up in a suburb outside of a large Midwestern city.</p> <p>This person has a dog, enjoys being physically active, follows college sports and frequently volunteers to work in local candidates' campaigns.</p>
Partisan (D)	<p>Suppose a new person moves into your neighborhood. The person is a registered Democrat, approximately 45 years old, has a college degree, and grew up in a suburb outside of a large Midwestern city.</p> <p>This person has a dog, enjoys being physically active, follows college sports and frequently volunteers to work in local candidates' campaigns.</p>
PID + ideology (R)	<p>Suppose a new person moves into your neighborhood. The person is a registered Republican, approximately 45 years old, has a college degree, and grew up in a suburb outside of a large Midwestern city.</p> <p>This person believes that humans and dinosaurs walked the earth at the same time, that elementary school students should be required to recite the pledge of allegiance, and that we should build a fence between the United States and Mexico.</p> <p>This person has a dog, enjoys being physically active, follows college sports and frequently volunteers to work in local candidates' campaigns.</p>
PID + ideology (D)	<p>Suppose a new person moves into your neighborhood. The person is a registered Democrat, approximately 45 years old, has a college degree, and grew up in a suburb outside of a large Midwestern city.</p> <p>This person believes that the country would be better if every citizen drove an electric car, that marijuana use should be legal in all states, and that a nationalized healthcare system would improve health for all citizens.</p> <p>This person has a dog, enjoys being physically active, follows college sports and frequently volunteers to work in local candidates' campaigns.</p>

Table A.3: Question Wording for Dependent Variables: TAPS Experiment

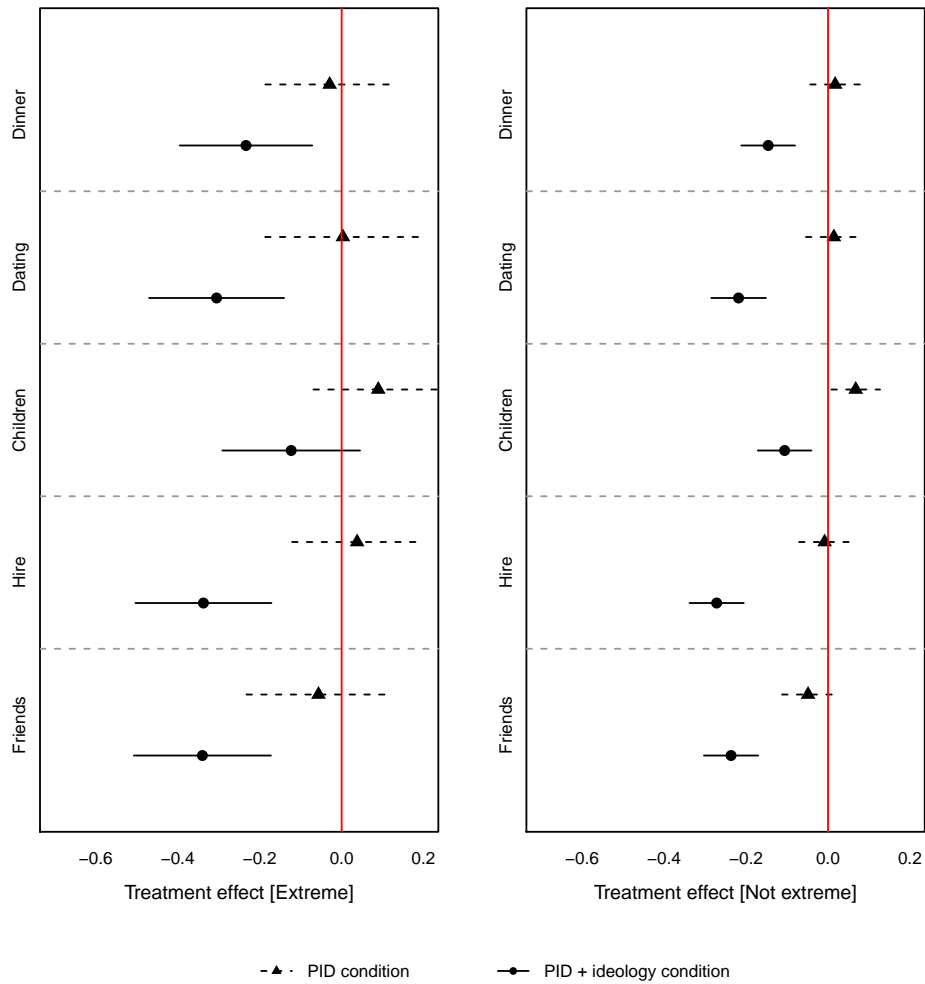
Short title	Full text
Friends	I would be interested in becoming friends with this person.
Hire	If I were an employer and needed to hire a new employee, I would be interested in considering this person for a position with my company.
Children	I would feel comfortable allowing my children to play with the new neighbor's kids.
Dating	I would consider dating this person or would introduce this person to a friend.
Dinner	I would consider having this person over for a family meal or barbecue.

Table A.4: Summary Statistics Across Experimental Treatments (TAPS)

Condition	Friends	Hire	Children	Dating	Dinner	Ideological Placement
Control	0.69	0.70	0.69	0.51	0.71	3.10
Partisan	0.64	0.69	0.75	0.52	0.71	3.04
PID + ideology	0.44	0.42	0.58	0.28	0.54	3.14

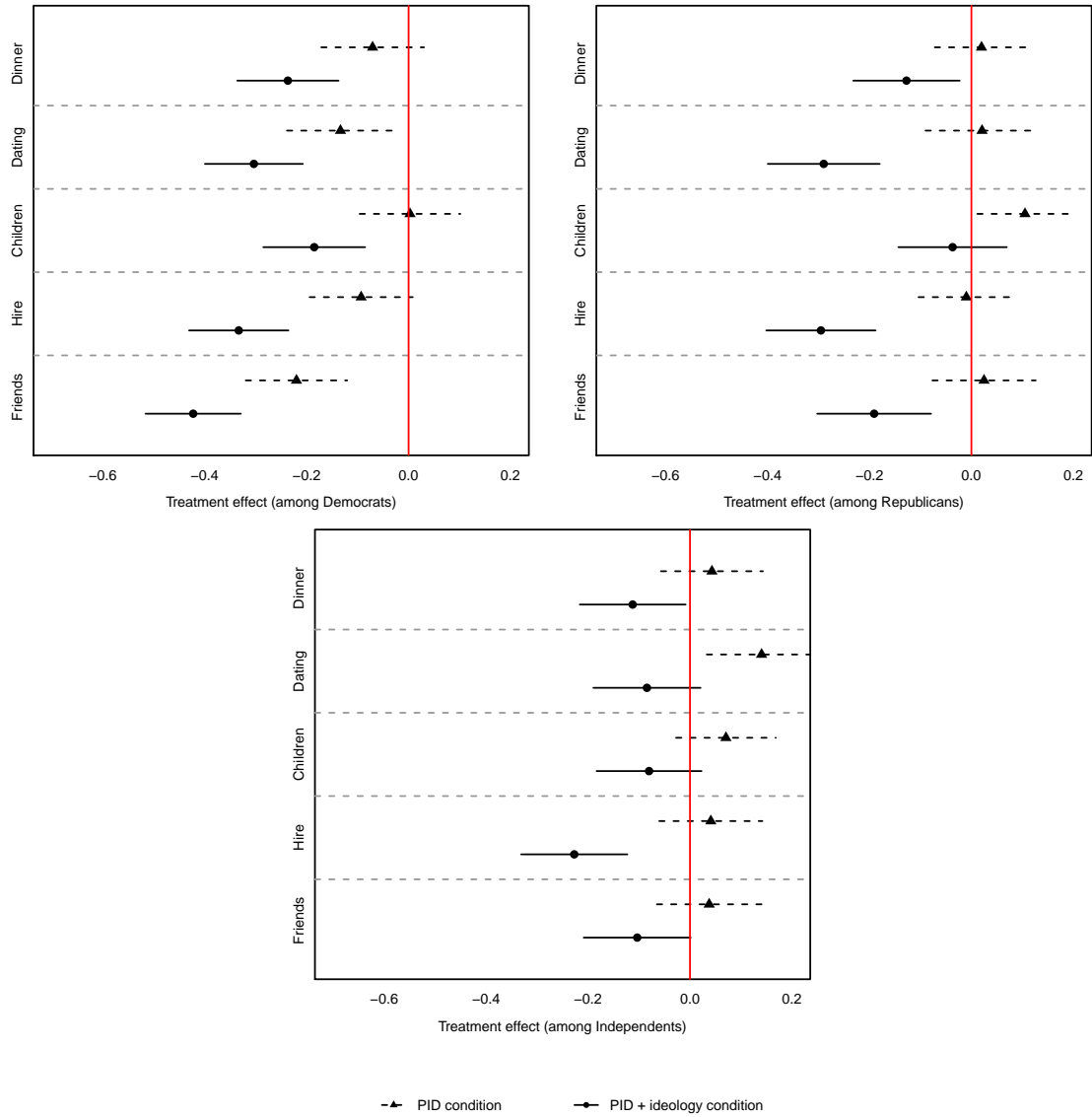
Entries are the proportion of respondents who agreed they would be interested in interacting with a potential neighbor in the ways indicated by the column headings. Entries in the last column reflect the mean placement on a five-point ideological scale ranging from very liberal (1) to very conservative (5). Sample sizes were N=491 for the control condition, N=497 for the partisan condition, and N=499 for the PID + ideology condition.

Figure A.1: Effect of Partisanship and Ideology on Interpersonal Evaluations by Ideological Extremity



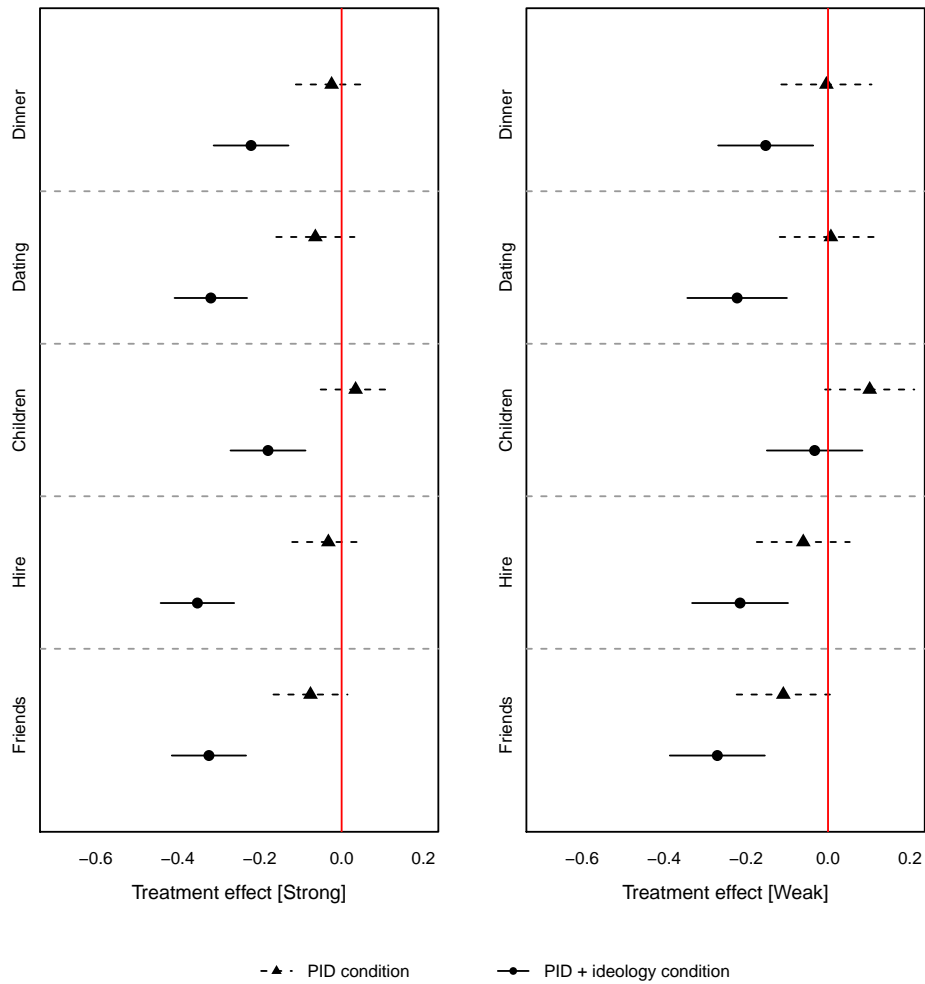
Note: Values along the x-axis indicate the difference in proportions when comparing each of the treatment groups to the control group. The vertical line at zero indicates the null hypothesis of no treatment effect. The horizontal lines show the 95% confidence intervals. The left plot shows individuals who are “very conservative” or “very liberal” and the right plot shows results for respondents who report more moderate ideologies.

Figure A.2: Effect of Partisanship and Ideology on Interpersonal Evaluations by PID



Note: Values along the x -axis indicate the difference in proportions when comparing each of the treatment groups to the control group. The vertical line at zero indicates the null hypothesis of no treatment effect. The horizontal lines show the 95% confidence intervals.

Figure A.3: Effect of Partisanship and Ideology on Interpersonal Evaluations by Strength of Partisan Identity



Note: Values along the x-axis indicate the difference in proportions when comparing each of the treatment groups to the control group. The vertical line at zero indicates the null hypothesis of no treatment effect. The horizontal lines show the 95% confidence intervals. The left plot shows individuals who identify as “strong” Republicans or Democrats and the right plot shows results for respondents who are identify as “weak” Republicans or Democrats.

Table A.5: Partisanship and the Heterogeneous Effect of Partisanship and Ideology on Social Interaction

	Friends	Hire	Children	Dating	Dinner
Partisan condition	-0.08 (0.05)	-0.06 (0.05)	0.03 (0.05)	-0.03 (0.05)	-0.07 (0.05)
PID + ideology condition	-0.31* (0.05)	-0.29* (0.05)	-0.13* (0.05)	-0.27* (0.05)	-0.18* (0.05)
Republican	0.01 (0.06)	-0.04 (0.05)	-0.05 (0.05)	-0.002 (0.06)	0.02 (0.05)
Partisan condition × Republican	-0.05 (0.08)	0.01 (0.08)	0.05 (0.07)	-0.06 (0.08)	0.09 (0.08)
PID + ideology condition × Republican	-0.05 (0.08)	-0.07 (0.08)	0.002 (0.07)	-0.07 (0.08)	-0.04 (0.08)
Constant	0.67* (0.12)	0.57* (0.12)	0.32* (0.12)	0.50* (0.12)	0.45* (0.12)
Controls	✓	✓	✓	✓	✓
Observations	929	929	929	929	929
R ²	0.09	0.11	0.07	0.10	0.07

Note: Table entry coefficients are estimated with Ordinary Least Squares. Standard errors are in parentheses. The dependent variable is listed at the top of each column. Control variables include *Female*, *White*, *Age*, *Education*, and *Income*. * $p < 0.05$

Table A.6: Strength of Partisanship and the Heterogeneous Effect of Partisanship and Ideology on Social Interaction

	Friends	Hire	Children	Dating	Dinner
Partisan condition	-0.11 (0.07)	-0.09 (0.07)	0.07 (0.07)	-0.01 (0.07)	-0.02 (0.07)
PID + ideology condition	-0.32* (0.07)	-0.23* (0.07)	-0.07 (0.07)	-0.24* (0.07)	-0.17* (0.07)
Strong Partisan	-0.06 (0.06)	-0.02 (0.06)	-0.03 (0.06)	-0.01 (0.06)	-0.03 (0.06)
Partisan condition × Strong Partisan	0.01 (0.08)	0.04 (0.08)	-0.01 (0.08)	-0.07 (0.08)	-0.01 (0.08)
PID + ideology condition × Strong Partisan	-0.03 (0.08)	-0.15 (0.08)	-0.09 (0.08)	-0.11 (0.08)	-0.05 (0.08)
Constant	0.68* (0.13)	0.54* (0.13)	0.31* (0.13)	0.47* (0.13)	0.45* (0.13)
Controls	✓	✓	✓	✓	✓
Observations	857	857	857	857	857
R ²	0.10	0.11	0.07	0.10	0.07

Note: Table entry coefficients are estimated with Ordinary Least Squares. Standard errors are in parentheses. The dependent variable is listed at the top of each column. Control variables include *Female*, *White*, *Age*, *Education*, and *Income*. * $p < 0.05$

B TAPS Study – Study 2, Perceptions

The following tables and figures provide additional background information regarding our panel, the measurement and distributions of key variables, and also the results from a series of robustness checks for the second part of our analysis that focuses on increased inaccuracy and exaggeration in perceptions of partisan out-groups as a result of ideological extremism. More specifically, Table B.1 provides descriptive statistics of the socio-demographic and political variables that define our sample, and Table B.2 compares the descriptive statistics of our two-wave sample to the statistics for each of the separate waves in order to make sure that our combined sample is still nationally representative. Table B.3 contains the full wording of the partisan perceptions questions. Table B.4 presents the detailed results of the factor analysis we ran to create the operational ideology scores, while Table B.5 lists the exact question wording for the items used. Figure B.1 displays the distribution the ideology scores we use, split up by partisan identification.

Table B.6 reports the full regression results for our main results table in the manuscript including the coefficient estimates for all control variables. Table B.7 uses the same modeling strategy, but examines panelists' perceived level of agreement with these statements among *in-party* members, and Table B.8 repeats the analysis for respondents who are Independents. The models in Table B.9 include an indicator for whether a respondent is a strong partisan to explore whether our results are driven by strength of party identification rather than ideological extremity (where we consider respondents coded as a "1" (for Democrats) or a "7" (for Republicans) on the 7-point party ID scale as strong partisans). Tables B.10 and B.11 substitute a traditional seven-point symbolic ideology variable for our measure of operational ideology. To disentangle ideological extremity from the strength of one's ideological identity, the specifications in Tables B.12 and B.13 control for both symbolic and operational ideology in the same models.

Table B.14 accounts for the partisan's own position on each statement to explore the possibility that respondents may hold exaggerated perceptions in ways that varied systematically with their position on that statement. For the results in Table B.15, we changed our measure of party identification to also classify *leaners* (people thinking of themselves as closer to one of the parties in a follow-up question) as partisans rather than as Independents. Table B.16 includes a *Sorted* indicator, which controls for the possibility that partisan ideological sorting might be driving our results, and the models in Table B.17 include the controls for strong partisanship, symbolic ideology, and partisan ideological sorting all in the same model specification. Because our outcome measure is a five-category variable, Table B.18 replicates the main analysis using ordered logit models. Tables B.19 and B.20 focus on one individual statement at a time instead of pooling them all together. Finally, Tables B.21 (all respondents) and B.22 (only original respondents) are based

on a follow-up survey in December 2016, where we asked respondents to indicate their answers on a 100-point scale to make sure that our answers are not driven by the five-category scheme.

All these robustness checks produce results that are consistent with the findings we report in the main manuscript.

Table B.1: Descriptive Statistics

Variable	Mean	Standard Deviation	Minimum	Maximum
Female	0.490	0.500	0	1
White	0.745	0.436	0	1
Age: 18-29	0.078	0.269	0	1
Age: 30-44	0.195	0.397	0	1
Age: 45-59	0.327	0.469	0	1
Age: 60+	0.400	0.490	0	1
Education: Less than High School	0.030	0.171	0	1
Education: High School/Some College	0.365	0.482	0	1
Education: College or more	0.605	0.489	0	1
Income: Less than \$30,000	0.204	0.403	0	1
Income: \$30,000-\$49,999	0.195	0.397	0	1
Income: \$50,000-\$79,999	0.261	0.439	0	1
Income: \$80,000+	0.340	0.474	0	1
Metropolitan	0.853	0.354	0	1
Region: Northeast	0.151	0.358	0	1
Region: Midwest	0.263	0.440	0	1
Region: South	0.356	0.479	0	1
West	0.231	0.421	0	1
Conservatism	-0.013	1.045	-1.907	2.340
Ideology	4.108	1.741	1	7
Democrat	0.413	0.493	0	1
Republican	0.248	0.432	0	1
Independent	0.339	0.474	0	1
Political Knowledge	6.948	2.306	0	10
Political interest: not at all interested	0.031	0.173	0	1
Political interest: not very interested	0.144	0.351	0	1
Political interest: somewhat interested	0.374	0.484	0	1
Political interest: very interested	0.451	0.498	0	1

Table B.2: Descriptive Statistics and Differences Between Waves

Variable	Difference		Difference
	Both waves	Both waves - Wave 1	Both waves - Wave 2
Female	0.49	0.03	0.01
White	0.74	-0.01	0.00
Age: 18-29	0.08	0.01	0.00
Age: 30-44	0.20	0.02	0.00
Age: 45-59	0.33	-0.01	0.00
Age: 60+	0.40	-0.02	0.00
Education: Less than High School	0.03	0.01	0.00
Education: High School/Some College	0.37	0.02	0.00
Education: College and more	0.60	-0.02	-0.01
Income: Less than \$30,000	0.20	0.02	0.01
Income: \$30,000-\$49,999	0.20	0.01	0.00
Income: \$50,000-\$79,999	0.26	-0.01	0.00
Income: More than \$80,000	0.34	-0.02	-0.01
Metropolitan	0.85	0.00	0.00
Region: Northeast	0.15	0.00	0.00
Region: Midwest	0.26	0.00	0.00
Region: South	0.36	0.01	0.00
Region: West	0.23	0.00	0.00
Conservatism	-0.01	0.01	0.01
Ideology	4.11	0.00	0.01
Democrat	0.41	-0.02	-0.01
Republican	0.25	0.00	0.00
Independent	0.34	0.02	0.01
Political Knowledge	6.95	-0.33*	-0.12
Political interest: not at all interested	0.03	0.01	0.00
Political interest: not very interested	0.14	0.01	0.00
Political interest: somewhat interested	0.37	0.01	0.00
Political interest: very interested	0.45	-0.04*	-0.01

Note: Column 1 shows the mean value of each variable (proportion of each demographic category) for respondents that completed the first *and* second wave of the survey. Column 2 presents the difference between the mean values of the sample that completed both waves and everybody who completed wave 1 (i.e., also including some respondents that did not complete wave 2). Similarly, Column 3 shows the differences between respondents in both waves and the full sample of wave 2. * indicates whether the differences are statistically significant at the 95% level or more (two-tailed t-tests). Respondents per wave: Wave 1 (March 2014) = 1,669; Wave 2 (May 2014) = 1,496; Both waves = 1,301.

Table B.3: Partisan Perceptions Questions

Republican perceptions	Democratic perceptions
This country would be safer if every law-abiding citizen possessed a firearm.	This country would be better if every citizen drove an electric car.
Humans and dinosaurs walked the earth at the same time.	Marijuana use should be legal in all states.
Homosexuality threatens the well-being of our country.	The federal government should impose a ban on the sale of soda.
Elementary students should be required to recite the pledge of allegiance every day.	This country would be better if we all paid more taxes.
We should build a fence between the United States and Mexico.	A nationalized healthcare system would improve health for all citizens.

These questions were preceded by the following instructions: “Please indicate whether you agree or disagree with each of the following statements.” Responses were measured on a five-point scale ranging from “strongly agree” to “strongly disagree,” with an additional “don’t know” response option. Respondents are coded as agreeing with the statement if they chose either the “strongly agree” or “somewhat agree” response options.

Table B.4: Operational Ideology Factor Analysis

Variable	Factor 1 Loading	Uniqueness
Abortion	0.63	0.45
Education Spending	-0.63	0.56
Taxes	0.72	0.43
Gay Marriage	0.62	0.44
Gun Control	0.57	0.63
Global Warming	0.80	0.35
ObamaCare	-0.76	0.38
Immigration	0.58	0.57
Minimum Wage	0.81	0.30
Medicaid Expansion	0.68	0.48
Regulation of Business	0.72	0.44
Affirmative Action	0.62	0.52
Privatization of Social Security	-0.42	0.75

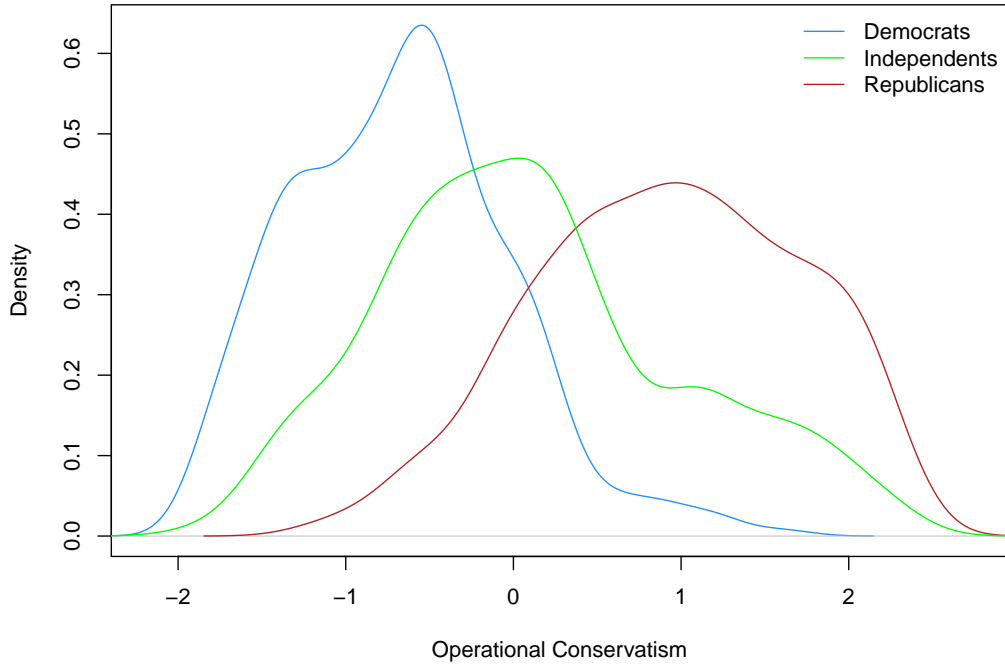
Note: First Factor Eigenvalue 5.77; Second Factor Eigenvalue 0.51; $\alpha = 0.91$

Table B.5: Question Wording for Operational Ideology Items

Item	Full text
Abortion	Federal programs that provide health care benefits should allow funding for abortions.
Education Spending	Federal spending for education should be reduced.
Taxes	Federal personal income taxes for individuals with incomes higher than \$250,000 should be raised.
Gay Marriage	The federal government should recognize the validity of a same-sex marriage where state law does.
Gun Control	Federal law should ban the possession of handguns except by law enforcement personnel.
Global Warming	The federal government should adopt policies to address the problem of global warming.
ObamaCare	The federal health care reform program adopted in 2010 should be repealed.
Immigration	The federal government should find a way to allow people who now are in the U.S. illegally to stay in the U.S. and become U.S. citizens.
Minimum Wage	The federal government should guarantee a higher minimum hourly wage for workers.
Medicaid Expansion	Medicaid, the federal government health program for low income people, should be extended to cover more people.
Regulation of Business	The federal government should do more to regulate business in order to protect the interests of consumers.
Affirmative Action	The federal government should support programs designed to help minorities better jobs and education.
Privatization of Social Security	Social Security should be reformed so that individuals are given private retirement accounts that are invested in the stock market.

Question Prompt: We are interested in your views on issues being debated in American national politics. For each issue, we give you a statement about the issue. We want to know whether you strongly agree, agree, neither agree nor disagree, disagree, or strongly disagree with each statement.

Figure B.1: Operational Ideology by Party



Note: The mean ideological score for Democrats is -0.68 , with a standard deviation of 0.64 . For Republicans, the mean score is 0.92 , with a standard deviation of 0.78 . For independents, the mean score is 0.10 , with a standard deviation of 0.90 .

Table B.6: Partisans' Views of Out-Party Members

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	0.336* (0.073)	-0.370* (0.069)	
Ideological Extremity			0.459* (0.066)
Political Knowledge	0.038 (0.029)	0.060* (0.022)	0.047* (0.018)
Political Interest	0.275* (0.077)	0.097 (0.056)	0.168* (0.045)
Income	-0.012 (0.015)	0.010 (0.013)	0.001 (0.010)
Education	-0.028 (0.035)	-0.025 (0.024)	-0.028 (0.019)
White	-0.018 (0.139)	0.011 (0.094)	-0.011 (0.077)
Female	0.054 (0.106)	0.061 (0.084)	0.061 (0.065)
Age 30-44	-0.297 (0.181)	-0.067 (0.169)	-0.109 (0.126)
Age 45-59	-0.302 (0.184)	-0.101 (0.156)	-0.139 (0.120)
Age 60+	-0.441* (0.183)	-0.151 (0.153)	-0.242* (0.118)
Metropolitan Area	-0.033 (0.132)	0.067 (0.117)	0.033 (0.088)
Midwest	-0.121 (0.170)	-0.144 (0.122)	-0.146 (0.098)
South	0.040 (0.155)	0.048 (0.120)	0.011 (0.094)
West	-0.095 (0.177)	0.206 (0.117)	0.078 (0.098)
Constant	2.224* (0.508)	2.530* (0.311)	2.005* (0.268)
Question FE	✓	✓	✓
<i>N</i> (Total)	1513	2480	3993
<i>N</i> (Respondents)	323	537	860
<i>R</i> ²	0.23	0.19	0.20

Table B.7: Partisans' Views of In-Party Members

	Views of Own Party		
	Republicans' Views of Republicans	Democrats' Views of Democrats	All Partisans' Views of In-Party Members
Conservatism	0.157* (0.069)	-0.216* (0.061)	
Ideological Extremity			0.268* (0.052)
Political Knowledge	-0.012 (0.025)	0.015 (0.017)	0.003 (0.014)
Political Interest	0.094 (0.069)	-0.059 (0.038)	-0.011 (0.035)
Constant	3.416* (0.434)	2.833* (0.258)	2.729* (0.219)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1487	2527	4014
<i>N</i> (Respondents)	323	547	870
<i>R</i> ²	0.26	0.29	0.29

Note: Table entries are linear regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is Perception_{*ij*}, the degree to which a respondent *i* believes members of the *in*-party agree with perception *j* as described in the text. Additional socio-demographic controls (income, education, sex, race, age, Metropolitan Statistical Area, and US Census region) and question indicators are included but not reported. * *p* < 0.05

Table B.8: Independents' Views of Partisans

	Views of Partisans	
	Views of Democrats	Views of Republicans
Conservatism	0.288* (0.047)	-0.127* (0.047)
Political Knowledge	0.058* (0.022)	0.029 (0.025)
Political Interest	-0.032 (0.058)	0.024 (0.067)
Constant	2.784* (0.377)	3.133* (0.428)
Controls	✓	✓
Question FE	✓	✓
<i>N</i> (Total)	2000	1999
<i>N</i> (Respondents)	441	439
<i>R</i> ²	0.20	0.18

Note: Table entries are linear regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is Perception_{ij} , the degree to which a respondent i believes Democrats/Republicans agree with item j as described in the text. Additional socio-demographic controls (income, education, sex, race, age, Metropolitan Statistical Area, and US Census region) and question indicators are included but not reported. * $p < 0.05$

Table B.9: Partisans' Views of Out-Party Members, Controlling for Strong Partisans

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	0.315* (0.081)	-0.384* (0.074)	
Ideological Extremity			0.451* (0.070)
Strong Partisan	0.099 (0.118)	-0.055 (0.091)	0.034 (0.071)
Political Knowledge	0.041 (0.029)	0.061* (0.022)	0.047* (0.018)
Political Interest	0.256* (0.079)	0.100 (0.056)	0.164* (0.046)
Constant	2.235* (0.506)	2.532* (0.310)	2.010* (0.268)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1513	2480	3993
<i>N</i> (Respondents)	323	537	860
<i>R</i> ²	0.23	0.19	0.20

Note: Table entries are ordinary least squares coefficients with standard errors in parentheses, clustered on individuals. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. * $p < 0.05$

Table B.10: Partisans' Views of Out-Party Members, Using Symbolic Conservatism

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
7-Point Symbolic Ideology (1=Very Lib., 7=Very Cons.)	0.170* (0.048)	-0.110* (0.032)	
Symbolic Ideological Extremity			0.144* (0.032)
Political Knowledge	0.056* (0.027)	0.084* (0.022)	0.079* (0.017)
Political Interest	0.252* (0.070)	0.124* (0.054)	0.187* (0.044)
Constant	1.496* (0.529)	2.775* (0.378)	1.770* (0.275)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1536	2418	3954
<i>N</i> (Respondents)	327	523	850
<i>R</i> ²	0.21	0.18	0.18

Note: Table entries are linear regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is Perception_{ij} , the degree to which a respondent i believes members of the out-party agree with perception j as described in the text. Additional socio-demographic controls (income, education, sex, race, age, Metropolitan Statistical Area, and US Census region) and question indicators are included but not reported.

* $p < 0.05$

Table B.11: Independents' Views of Partisans, Using Symbolic Conservatism

	Views of Partisans	
	Views of Democrats	Views of Republicans
7-Point Sym. Ideol. (1=Very Lib., 7=Very Cons.)	0.173* (0.029)	-0.027 (0.032)
Political Knowledge	0.058* (0.023)	0.028 (0.026)
Political Interest	0.007 (0.060)	-0.011 (0.068)
Constant	2.049* (0.404)	3.275* (0.439)
Controls	✓	✓
Question FE	✓	✓
N (Total)	1955	1936
N (Respondents)	428	426
R^2	0.21	0.18

Note: Table entries are linear regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is Perception_{*i**j*} the degree to which a respondent *i* believes Democrats/Republicans agree with perception *j* as described in the text. Additional socio-demographic controls (income, education, sex, race, age, Metropolitan Statistical Area, and US Census region) and question indicators are included but not reported. * $p < 0.05$

Table B.12: Partisans' Views of Out-Party Members, Using Symbolic Conservatism and Operational Conservatism

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	0.322* (0.087)	-0.271* (0.077)	
Ideological Extremity			0.405* (0.072)
7-Point Symbolic Ideology (1=Very Lib., 7=Very Cons.)	0.079 (0.060)	-0.063 (0.035)	
Symbolic Ideological Extremity			0.076* (0.037)
Political Knowledge	0.040 (0.029)	0.062* (0.023)	0.047* (0.018)
Political Interest	0.211* (0.072)	0.070 (0.056)	0.137* (0.044)
Constant	2.048* (0.544)	2.780* (0.385)	1.994* (0.280)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1471	2337	3808
<i>N</i> (Respondents)	312	505	817
<i>R</i> ²	0.24	0.19	0.20

Note: Table entries are linear regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is Perception_{ij}, the degree to which a respondent *i* believes members of the out-party agree with perception *j* as described in the text. Additional socio-demographic controls (income, education, sex, race, age, Metropolitan Statistical Area, and US Census region) and question indicators are included but not reported.

* $p < 0.05$

Table B.13: Independents' Views of Partisans, Using Symbolic Conservatism and Operational Conservatism

	Views of Partisans	
	Views of Democrats	Views of Republicans
Conservatism	0.224* (0.060)	-0.218* (0.064)
7-Point Sym. Ideol. (1=Very Lib., 7=Very Cons.)	0.076* (0.038)	0.065 (0.043)
Political Knowledge	0.054* (0.023)	0.040 (0.026)
Political Interest	-0.021 (0.060)	-0.006 (0.069)
Constant	2.417* (0.427)	2.944* (0.445)
Controls	✓	✓
Question FE	✓	✓
<i>N</i> (Total)	1867	1855
<i>N</i> (Respondents)	409	408
<i>R</i> ²	0.22	0.20

Note: Table entries are linear regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is Perception_{*i**j*} the degree to which a respondent *i* believes Democrats/Republicans agree with perception *j* as described in the text. Additional socio-demographic controls (income, education, sex, race, age, Metropolitan Statistical Area, and US Census region) and question indicators are included but not reported. * $p < 0.05$

Table B.14: Partisans' Views of Out-Party Members, Controlling for Agreement with Statement

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	0.338* (0.073)	-0.371* (0.069)	
Ideological Extremity			0.459* (0.066)
Agree with Statement	0.362 (0.305)	0.019 (0.097)	-0.003 (0.094)
Political Knowledge	0.039 (0.029)	0.060* (0.022)	0.047* (0.018)
Political Interest	0.276* (0.077)	0.097 (0.056)	0.168* (0.045)
Constant	2.194* (0.511)	2.524* (0.315)	2.006* (0.272)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1513	2480	3993
<i>N</i> (Respondents)	323	537	860
<i>R</i> ²	0.23	0.19	0.20

Note: Table entries are linear regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is Perception_{ij} , the degree to which a respondent i believes members of the out-party agree with perception j as described in the text. *Agree with Statement* is a dichotomous variable that indicates whether respondent i indicated agreement with perception j themselves. Additional socio-demographic controls (income, education, sex, race, age, Metropolitan Statistical Area, and US Census region) and question indicators are included but not reported. * $p < 0.05$

Table B.15: Partisans' Views of Out-Party Members (Including *Leaners*)

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	0.201* (0.045)	-0.287* (0.051)	
Ideological Extremity			0.375* (0.039)
Strength of Partisanship	-0.025 (0.036)	-0.036 (0.033)	-0.048* (0.024)
Political Knowledge	0.017 (0.017)	0.036* (0.014)	0.023* (0.011)
Political Interest	0.096* (0.049)	0.002 (0.032)	0.038 (0.027)
Constant	3.119* (0.314)	2.363* (0.217)	2.621* (0.176)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	4957	6880	11837
<i>N</i> (Respondents)	548	772	1320
<i>R</i> ²	0.21	0.25	0.21

Note: Table entries are ordinary least squares coefficients with standard errors in parentheses, clustered on individuals. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. * $p < 0.05$

Table B.16: Partisans' Views of Out-Party Members, Controlling for Sorting

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	0.395* (0.095)	-0.313* (0.094)	
Ideological Extremity			0.415* (0.068)
Sorted	-0.233 (0.188)	0.153 (0.174)	0.262* (0.099)
Political Knowledge	0.037 (0.028)	0.062* (0.022)	0.046* (0.018)
Political Interest	0.267* (0.078)	0.097 (0.056)	0.163* (0.044)
Constant	2.379* (0.522)	2.423* (0.349)	1.872* (0.271)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1513	2480	3993
<i>N</i> (Respondents)	323	537	860
<i>R</i> ²	0.23	0.19	0.20

Note: Table entries are ordinary least squares coefficients with standard errors in parentheses, clustered on individuals. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. *Sorted* is coded as 1 if a panelist identifies as a partisan and lies on the “sorted” side of the Conservatism scale. Likewise, a panelist is coded as 0 if they are a partisan and remain on the “unsorted” side of the conservatism scale. For example, if a Democrat scores less than the standardized mean of 0, they are coded as 1. Similarly, if a Republican scores less than 0, they are coded as 0. * $p < 0.05$

Table B.17: Partisans' Views of Out-Party Members, Using Symbolic Conservatism, Strong Partisanship, Sorting, and Operational Conservatism

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	0.403* (0.105)	-0.225* (0.104)	
Ideological Extremity			0.375* (0.074)
Strong Partisan	0.034 (0.118)	-0.032 (0.092)	0.008 (0.072)
7-Point Symbolic Ideology (1=Very Lib., 7=Very Cons.)	0.077 (0.062)	-0.063 (0.035)	
Symbolic Ideological Extremity			0.067 (0.037)
Sorted	-0.360 (0.185)	0.155 (0.174)	0.226* (0.101)
Political Knowledge	0.039 (0.028)	0.065* (0.023)	0.046* (0.018)
Political Interest	0.191* (0.075)	0.072 (0.056)	0.132* (0.046)
Constant	2.317* (0.555)	2.678* (0.418)	1.891* (0.282)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1471	2337	3808
<i>N</i> (Respondents)	312	505	817
<i>R</i> ²	0.24	0.19	0.20

Note: Table entries are ordinary least squares coefficients with standard errors in parentheses, clustered on individuals. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. * $p < 0.05$

Table B.18: Partisans' Views of Out-Party Members, Ordered Logit

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	0.545* (0.117)	-0.543* (0.102)	
Ideological Extremity			0.719* (0.101)
Political Knowledge	0.052 (0.044)	0.079* (0.034)	0.058* (0.027)
Political Interest	0.408* (0.121)	0.137 (0.084)	0.248* (0.067)
21 to 40 percent	-0.656 (0.819)	-0.949 (0.465)	-0.260 (0.401)
41 to 60 percent	0.554 (0.813)	0.000 (0.461)	0.787 (0.397)
61 to 80 percent	1.716 (0.812)	1.064 (0.464)	1.884 (0.399)
81 to 100 percent	3.054 (0.824)	2.480 (0.471)	3.267 (0.406)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1513	2480	3993
<i>N</i> (Respondents)	323	537	860
log likelihood	-2218.50	-3688.85	-5922.66
Wald χ^2	311.18	360.51	653.36

Note: Table entries are ordered logit regression coefficients with standard errors in parentheses, clustered on individuals. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. * $p < 0.05$

Table B.19: Individual Questions

Republicans' Views of Democrats					
	Electric Car	Marijuana	Soda Ban	Taxes	Health Care
Conservatism	0.297* (0.107)	0.325* (0.098)	0.405* (0.102)	0.569* (0.104)	0.092 (0.091)
Political Knowledge	0.061 (0.042)	0.058 (0.040)	0.008 (0.042)	0.051 (0.042)	0.023 (0.038)
Political Interest	0.313* (0.107)	0.325* (0.100)	0.266* (0.111)	0.109 (0.104)	0.326* (0.094)
Constant	2.659* (0.699)	1.577* (0.657)	3.083* (0.692)	2.348* (0.679)	2.387* (0.626)
Controls	✓	✓	✓	✓	✓
<i>N</i>	291	302	293	316	311
<i>R</i> ²	0.15	0.15	0.15	0.20	0.09

Note: Table entries are ordinary least squares coefficients with standard errors in parentheses. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. * $p < 0.05$

Table B.20: Individual Questions

Democrats' Views of Republicans					
	Guns	Dinosaurs	Homosexuality	Pledge	Border Fence
Conservatism	-0.463*	-0.121	-0.468*	-0.356*	-0.410*
	(0.101)	(0.116)	(0.100)	(0.108)	(0.094)
Political Knowledge	0.111*	-0.032	0.052	0.100*	0.059
	(0.030)	(0.035)	(0.032)	(0.033)	(0.029)
Political Interest	0.058	0.202*	0.114	0.031	0.078
	(0.073)	(0.087)	(0.079)	(0.079)	(0.072)
Constant	2.172*	1.522*	2.149*	2.838*	3.009*
	(0.426)	(0.511)	(0.444)	(0.476)	(0.420)
Controls	✓	✓	✓	✓	✓
<i>N</i>	519	463	496	496	506
<i>R</i> ²	0.16	0.06	0.13	0.14	0.11

Note: Table entries are ordinary least squares coefficients with standard errors in parentheses. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. * $p < 0.05$

Table B.21: Partisans' Views of Out-Party Members Using 100 Point Scale – *All respondents*

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	9.440* (1.661)	-6.694* (1.867)	
Ideological Extremity			9.952* (1.535)
Political Knowledge	1.753* (0.695)	0.631 (0.603)	0.866 (0.461)
Political Interest	0.210 (1.539)	-0.989 (1.411)	-0.158 (1.011)
Constant	53.312* (9.769)	62.760* (8.320)	51.106* (6.290)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1539	2249	3788
<i>N</i> (Respondents)	307	454	761
<i>R</i> ²	0.20	0.17	0.18

Note: Table entries are ordinary least squares coefficients with standard errors in parentheses, clustered on individuals. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. * $p < 0.05$

Table B.22: Partisans' Views of Out-Party Members Using 100 Point Scale – *Only original respondents*

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	8.759* (1.699)	-5.811* (1.943)	
Ideological Extremity			8.868* (1.540)
Political Knowledge	1.725* (0.736)	0.812 (0.610)	0.984* (0.469)
Political Interest	-0.009 (1.625)	-0.654 (1.527)	-0.130 (1.078)
Constant	57.517* (10.704)	57.707* (8.735)	50.065* (6.712)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1307	1951	3258
<i>N</i> (Respondents)	280	418	698
<i>R</i> ²	0.19	0.18	0.18

Note: Table entries are ordinary least squares coefficients with standard errors in parentheses, clustered on individuals. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. * $p < 0.05$

C TAPS Study – Correct Assessment

As part of our analysis of (mis)perceptions, we examine how well respondents identify the correct proportion of out-party members who agree with a given statement. For each item, we created a measure of *correct assessment* that indicates whether a respondent correctly identified the percentage category that contained the observed level of agreement with the partisan perception. For example, 28.4% of all Republicans agreed with the statement that “Humans and dinosaurs walked the earth at the same time.” A Democratic respondent who is asked about the percentage of Republicans they believed agreed with that statement are coded as “1” for *correct assessment* if they believed that 21-40% of Republicans agreed with this statement, while Democrats that believed that 0-20, 41-60, 61-80, or 81-100% of Republicans agree with the statement are coded as “0” for *correct assessment*.

We model this variable in a series of logit models that include our measure of ideology, item-specific indicators, and standard errors clustered on respondents:

$$\Pr(\text{CorrectAssessment}_{ij} = 1) = \text{logit}^{-1}(\beta_1 + \beta_2 \text{Ideology}_i + \gamma \mathbf{X}_i + \delta_j + \varepsilon_{ij}),$$

The results can be found in Table C.1. In short, we find that the likelihood of identifying the “correct proportion” of opposite party supporters who believe a agree with a given statement is significantly related to the respondent’s own ideology. For example, the more conservative a Republican is, the less likely they are to identify the correct proportion of Democrats that maintain an extreme view. Likewise, the more liberal a Democrat is, the less likely they are to provide the correct percentage of Republicans believing a given position. As for the shown control variables, we find limited evidence that political sophistication has a positive association with the ability to classify the opposite party. Although the variable had a positive coefficient estimate in the previous models, suggesting that more politically knowledgeable respondents believed partisans were more extreme, in Column 2 of Table C.1 we find that more sophisticated Democrats are significantly more likely to correctly classify Republicans.

To get a better idea of the substantive effect of our ideology measure in Table C.1, consider the following probabilities. Fixing all control variables at their respective mean, a Republican with an ideology score of -0.5 would be predicted to assess Democrats’ level of agreement with their perceptions correctly 36% of the time, whereas this would decrease to 23% at an an ideology score of 1.5. Similarly, for a very liberal Democrat with an ideology score of -1.5 , the model would expect them to correctly assess the level to which Republicans agree with perceptions of Republicans 19% of the time, whereas this would increase to 27% for an an ideology score of 0.25 while keeping all other variables at their mean. This in turn suggests that the effects we uncover

are not only statistically, but also substantively significant.

Table C.2 presents the results of a similar model that furthermore accounts for the possibility of measurement error, which could be induced by our five-category scheme as some of the actual rates of agreement are close to the line that divides one category from the next. In order to do so, we created a second version of the $\text{CorrectAssessment}_{ij}$ variable that also considers assessments as correct if they are within five percentage points of the correct category. The results are robust with the findings presented in Table C.1.

Figure C.1 provides additional information on the descriptive statistics regarding (mis)perceptions of counterpartisans' beliefs. The left panel of the figure shows second-order beliefs for Democratic items and the right panel shows second-order beliefs for Republican items. For each item we plot three bar charts that illustrate the degree to which Democrats, Independents, and Republicans believed that Democrats (in the left panel) and Republicans (in the right panel) agreed with their respective item, where the width of each colored region corresponds to the share of respondents that answered with that respective category. The category with the black frame reflects the *observed* answer (the correct level of agreement based on the figures shown in Table 1 in the main text) and the shaded category indicates the modal second-order belief.

For example, we know from Table 1 that 63.2% of Democrats agreed with the nationalized health care item (a stereotypically Democratic item). Examining the three bar charts for that item in Figure C.1, the fourth category (61-80%) has a bold frame, indicating that it contains the *correct* (observed) level of agreement. The shaded regions indicate that most Democrats and most Independents correctly identified that the observed level of agreement fell into this category. However, among Republicans the shaded region indicates that most of them believed that 81-100% of Democrats agreed with this statement. In other words, the modal Republican perceived an exaggerated level of agreement among Democrats for this item.

The right panel of Figure C.1 presents a similar pattern for the Republican items. Democrats exaggerated levels of agreement among Republicans, whereas Republicans themselves tended to correctly or slightly underestimate the level of agreement. For example, 42.5% of Republicans agreed that “this country would be safer if every law-abiding citizen possessed a firearm.” However, most Democrats and Independents believed the level of agreement among Republicans was 61-80%, while most Republicans correctly chose the 41-60% category.

For eight of our ten items, supporters of one party overestimated rates of agreement among members of the out-party for their respective item. More specifically, for four out of the five Republican items, Democrats exaggerated levels of agreement among Republicans, whereas Republicans similarly exaggerated levels of agreement among Democrats for four out of the

five Democratic items. Figure C.1 thus provides support for two conclusions. First, partisan respondents tend to exaggerate out-party perceptions, as Democratic (Republican) identifiers held partisan perceptions that were less accurate for Republicans (Democrats) than for Democrats (Republicans). Second, Independents demonstrated varied success in identifying partisans' levels of agreement with the statements. While they accurately categorized Democrats' beliefs for most items, they overestimated the perceived level of agreement for Republican items on all but one item.

Table C.1: Predicting Correct Assessment of Partisan Perceptions

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	-0.332* (0.091)	0.272* (0.095)	
Ideological Extremity			-0.312* (0.086)
Political Knowledge	-0.016 (0.036)	0.088* (0.035)	0.048 (0.025)
Political Interest	-0.100 (0.096)	-0.030 (0.085)	-0.056 (0.062)
Constant	-1.373* (0.637)	-2.469* (0.476)	-2.102* (0.384)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1513	2480	3993
<i>N</i> (Respondents)	323	537	860
Wald χ^2	82.39	56.35	121.27
log likelihood	-811.99	-1276.23	-2101.46

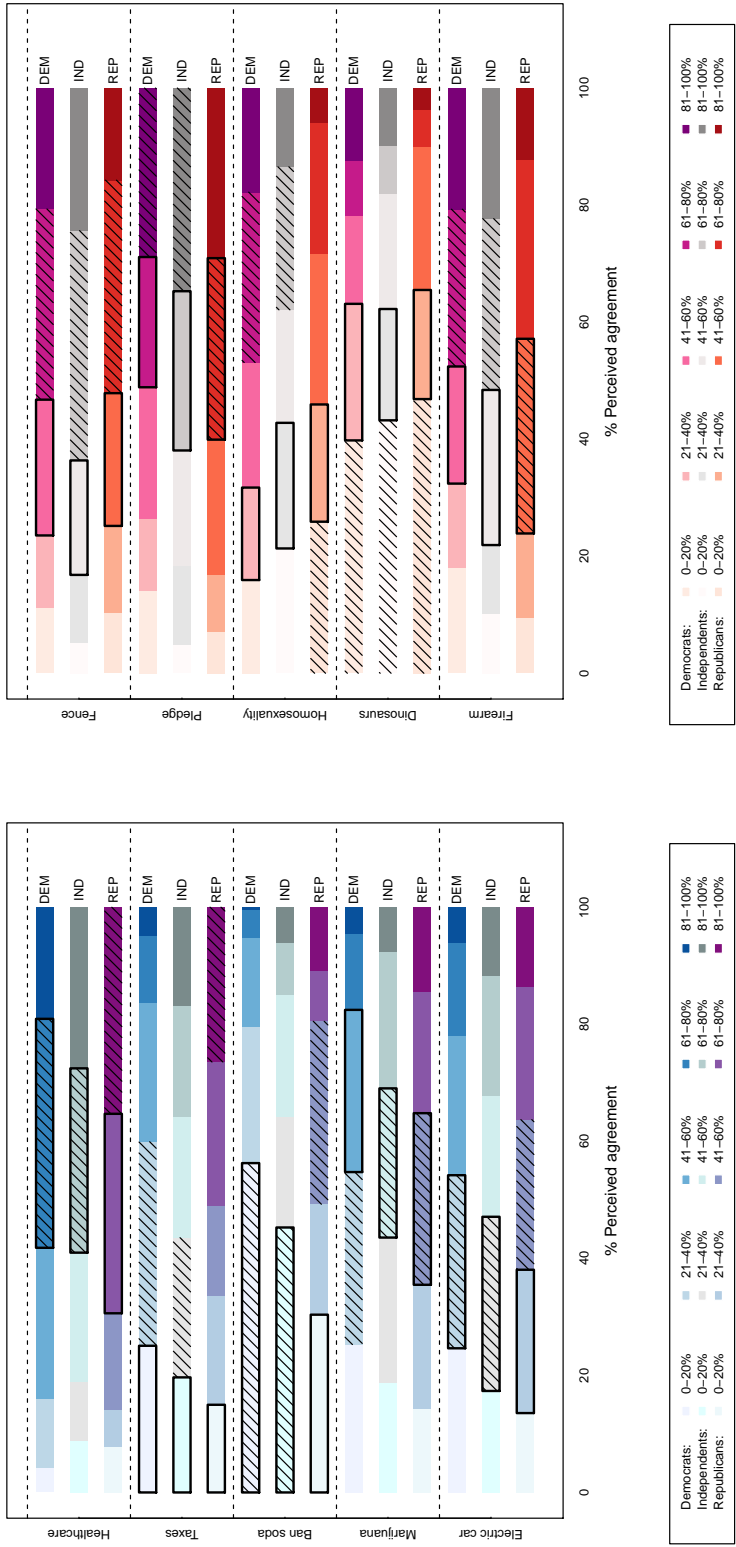
Note: Table entries are logit regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is $\text{CorrectAssessment}_{ij}$, indicating whether respondent i perceived the correct level of out-party agreement with perception j as described in the text. Additional socio-demographic controls (income, education, sex, race, age, Metropolitan Statistical Area, and US Census region) and question indicators are included but not reported. * $p < 0.05$

Table C.2: Ability to Determine Correct Percentage within Five Percent

	Views of Opposing Party		
	Republicans' Views of Democrats	Democrats' Views of Republicans	All Partisans' Views of Out-Party Members
Conservatism	-0.514* (0.101)	0.109 (0.106)	
Ideological Extremity			-0.374* (0.089)
Political Knowledge	0.010 (0.041)	0.127* (0.036)	0.078* (0.026)
Political Interest	-0.115 (0.097)	-0.057 (0.085)	-0.063 (0.063)
Constant	0.442 (0.699)	0.608 (0.478)	-.887* (0.412)
Controls	✓	✓	✓
Question FE	✓	✓	✓
<i>N</i> (Total)	1513	2480	3993
<i>N</i> (Respondents)	323	537	860
Wald χ^2	79.78	546.55	613.56
log likelihood	-962.37	-1185.14	-2177.60

Note: Table entries are logit regression coefficients with standard errors in parentheses, clustered on individuals. The outcome variable is $\text{CorrectAssessment}_{ij}$, indicating whether respondent i perceived the correct level of out-party agreement with perception j as described in the text, or was within five percentage points of that category. Additional socio-demographic controls and question dummy variables are included, but not reported. These controls include income, education, sex, race, age, Metropolitan Statistical Area, and US Census region. * $p < 0.05$

Figure C.1: Descriptive Statistics: Second-Order Beliefs



(a) Democrat Items

(b) Republican Items

Note: The bars with the **black frames** represent the category in which the *true* level of agreement among Democrats (left panel) and Republicans (right panel) with each statement shown in the y axis fall. Each *shaded* bar represents the modal response given by each partisan group regarding the perceived level of agreement of the group under analysis (Democrats, left panel; Republicans, right panel).